Lab 1: NMMAPS

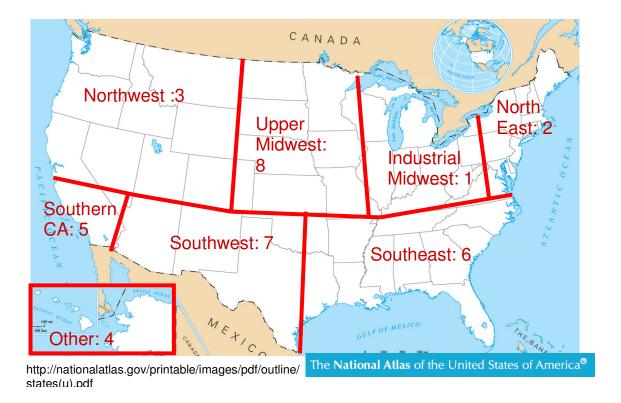
Data: City-specific air pollution effect estimates from an analysis of a 100 city subset of the National Morbidity Mortality and Air Pollution Study (NMMAPS) and city-level covariates. NMMAPS is a multi-city study containing city-specific daily time-series data on air pollution levels (including particulate matter of 10 microns and less in aerodynamic diameter – PM_{10}), mortality counts, temperature and dewpoint temperature.

Variables

- city: four letter city name abbreviation
- cityname: human readable name of the city
- state: state in which the city is located
- beta: city-specific coefficient on lag 1 PM₁₀ from a log-linear model relating changes in daily particulate matter air pollution to changes in daily mortality
- sebeta: standard error of beta
- population2000: city's population from 2000 census
- pdrive: proportion of the population that drives to work
- punem: proportion of the population unemployed
- pdeg: proportion of the population with a college degree or higher
- p65p: proportion of the city's population aged 65 or older
- latitude
- longitude
- altitude (contains some missing values)
- region: one of 7 NMMAPS regions (Industrial Midwest = IM, North East = NE, North West = NW, Southern California = SC, South East = SE, South West = SW, Upper Midwest = UM)

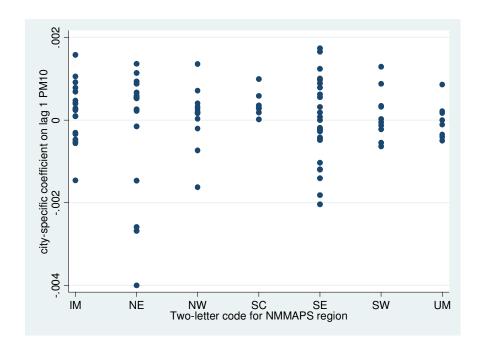
Goal: Estimate regional average associations between daily variations in PM₁₀ and daily variations in city-level mortality counts by combining city-specific estimates of log relative risks using shrinkage.

For now, we will ignore the standard error of the betas and consider only the variables beta and region.



Exploratory Data Analysis

- . sort region
- . scatter beta region, xlabel(1 2 3 4 5 6 7, valuelabel)



Approach A: No shrinkageCalculate each region's observed average coefficient on PM

 $\overline{\beta}_j = \frac{1}{n_j} \sum_{i=1}^{n_j} \beta_{ij}$ where β_{ij} is the city-specific estimate and j indexes region while i indexes city within region and n_j is the number of cities in region j.

. mean beta, over(region)

Mean estimation

IM: region = IM

NE: region = NE

NW: region = NW

SC: region = SC

SE: region = SE

SW: region = SW

UM: region = UM

	Over	Mean	Std. Err.	[95% Conf.	<pre>Interval]</pre>
		+			
beta					
	IM	.0001913	.0001602	0001265	.0005092
	NE	000156	.0003765	0009031	.0005912
	NW	.0001014	.0001951	0002858	.0004886
	SC	.0003898	.0001203	.0001512	.0006285
	SE	0000337	.0001989	0004283	.0003609
	SW	.000129	.0001911	0002501	.0005081
	UM	0000138	.0001564	0003241	.0002965

Approach B: Complete shrinkage

Calculate overall average of city-specific estimates

$$\overline{\beta} = \frac{\sum_{j=1}^{J} \frac{n_j}{\sigma^2} \overline{\beta}_j}{\sum_{j=1}^{J} \frac{n_j}{\sigma^2}} = \frac{\sum_{j=1}^{J} n_j \overline{\beta}_j}{\sum_{j=1}^{J} n_j} = \frac{\sum_{j=1}^{J} n_j \left(\frac{1}{n_j} \sum_{i=1}^{n_j} \beta_{ij}\right)}{n} = \frac{\sum_{j=1}^{J} \left(\sum_{i=1}^{n_j} \beta_{ij}\right)}{n} = \frac{\sum_{j=1}^{J} \beta_{ij}}{n}$$

. mean beta

Mean estimation Number of obs = 100

Which Approach should we use...A or B? Try an analysis of variance:

. oneway beta region

Analysis of Variance								
Source	SS	df	MS	F	Prob > F			
Between groups Within groups	2.2192e-06 .000084249	6 93	3.6987e-07 9.0591e-07	0.41	0.8719			
Total	.000086469	99	8.7342e-07					
Bartlett's test for	equal varianc	es:	chi2(6) = 30.5	327 Prok	o>chi2 = 0.000			

No evidence for difference in means of the regions since the F-statistic comparing the ratio of the MS_{between} to the MS_{within} has a p-value of 0.87. According to the ANOVA we should use approach B. But...the ANOVA requires the assumption that the variance within region is the same for each region. We have shown this to be false with Bartlett's test for equal variances. Hence, we will try approach C a compromise that uses a weighted combination of approaches A and B.

Approach C: Weighted combination of A and B

Taking a short cut in stata, we specify a model with a random intercept for region, then obtain the empirical Bayes estimates for each region

There are three ways we can specify a random intercept for our continuous outcome

```
xtreg - doesn't work for our data since likelihood too difficult to maximize
. xtreg beta, re i(region) mle
```

xtmixed – equiv to xtreg and doesn't work for same reason

```
. xtmixed beta || region:, mle
```

gllamm - works!

```
. gllamm beta, i(region) adapt nip(15)
Running adaptive quadrature
Iteration 0: log likelihood = 504.47871
Iteration 1: log likelihood = 507.63647
Iteration 2: log likelihood = 516.66851
Iteration 3: log likelihood = 526.03579
Iteration 4: log likelihood = 535.76318
Iteration 5: log likelihood = 545.78396
Iteration 6: log likelihood = 554.51778
Iteration 7: log likelihood = 554.60229
Iteration 8: log likelihood = 554.78871
Iteration 12:
               log likelihood = 556.15105
Adaptive quadrature has converged, running Newton-Raphson
Iteration 0: log likelihood = 556.15105
Iteration 1: log likelihood = 556.15105
number of level 1 units = 100
number of level 2 units = 7
Condition Number = 760.4234
```

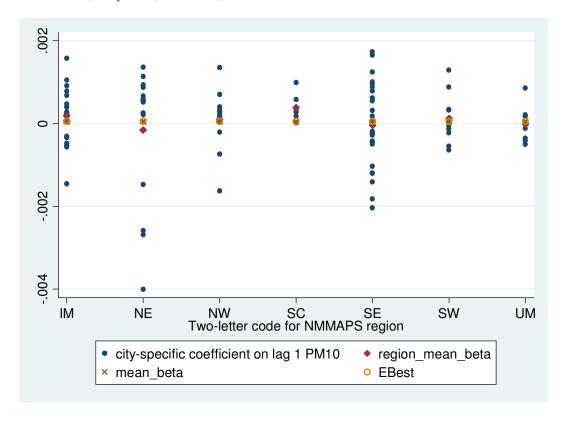
gllamm model

log likelihood = 556.15105

iog iinciinood	330.13103						
beta					[95% Conf.		
					0001291		
Variance at level 1							
8.650e-07 (1.224e-07)							
Variances and covariances of random effects							
***level 2 (reg	ion)						
var(1): 3.210e-16 (4.177e-12)							

We interpret 8.647e-07 to be the estimated variance of the residuals and 3.210e-16 to be the estimated variance of the region-specific intercepts.

scatter beta region_mean_beta mean_beta EBest region, xlabel(1 2 3 4 5 6 7,
valuelabel) msymbol(o d X Oh)



Note that the EB estimates are weighted combinations of the region_mean_beta (Approach A) and the mean_beta (Approach B). Most of the EB estimates fall closer to the region_mean_beta except for the NE region, which falls closer to the mean_beta estimate.