

Inference in Randomized Trials with Death and Missingness

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Outline

- 1 Motivation
- 2 Method
- 3 Software
- 4 Analysis results
- 5 Summary

Motivating example

- Randomized, double-blind, placebo-controlled phase III study
- Intent-to-treat population: **advanced** non-small cell lung **cancer** subjects
- Functional outcomes scheduled to be measured at baseline, 6 weeks and 12 weeks

Death and missingness

	Arm A	Arm B
	<i>n = 157</i>	<i>n = 322</i>
Died Prior to Wk 12	15%	17%
Survivors with complete data	59%	57%
Survivors missing only Wk 6	2%	5%
Survivors missing only Wk 12	11%	10%
Survivors missing both Wk 6 and 12	13%	11%

overall: 16% deaths; 30% survivors with missing data

Data truncated by death

Common analysis methods:

- Evaluate treatment effects **conditional** on survival
- **Joint** modeling survival and functional outcome
- Evaluate **causal** treatment effects for principal stratum
- **Composite** endpoint combining survival and functional outcomes

Goal

To propose a **composite outcome approach** that handles missing clinical evaluation data among subjects alive at the assessment times.

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General setting

- Consider a two arm randomized study with $T = 0, 1$
- Outcomes Y_0, \dots, Y_K collected at t_0, \dots, t_K , respectively
- Functional endpoint defined by $Z = f(Y_0, \dots, Y_K)$
 - example: $Z = Y_K$
 - example: $Z = Y_K - Y_0$
 - motivating study: $Z = (Y_2 + Y_1)/2 - Y_0$
- Survival time denoted by L
- Baseline covariates denoted by X
- Life status at t_K denoted by $\delta = I(L > t_K)$
- **Composite** endpoint: $C(L, \delta Z)$

Ranking

- Assume that **higher** values of Z denote **better** outcomes
- Assume no missing data at this moment
- Consider two subjects i and j with composite endpoint C_i and C_j , respectively
- $C_i > C_j$ (i **better** than j) only if
 - $\delta_i = \delta_j = 1$ and $Z_i > Z_j$, or
 - $\delta_i = \delta_j = 0$ and $L_i > L_j$, or
 - $\delta_i > \delta_j$
- Ranking may incorporate **clinical meaningful differences** in Z and L

Hypothesis testing

- Consider observing $C_{i,0}$ from subject i with $T = 0$, $C_{j,1}$ from subject j with $T = 1$
- Parameter of interest: $\theta = P(C_{i,0} > C_{j,1}) - P(C_{i,0} < C_{j,1})$
- $\theta = 0$ if no treatment effect
- Hypothesis: $H_0 : \theta = 0$ vs. $H_0 : \theta \neq 0$

Hypothesis testing

- Estimate θ by

$$\hat{\theta} = \frac{1}{n_0 n_1} \sum_{i: T_i=0} \sum_{j: T_j=1} \{I(C_i < C_j) - I(C_i > C_j)\}$$

- Variance of $\hat{\theta}$ available in closed form
- Consider the Wald test

Treatment effect size

- θ quantifies treatment effect size
- Recommend to compare quantiles (e.g. median) of $C(L, \delta Z)$ from each arm

Missingness

- For survivors ($\delta = 1$)
 - Denote τ_k to be the missingness indicator of Y_k
 - Denote $\mathbf{S} = (\tau_1, \dots, \tau_K)$ to be the missing pattern
- Intermittent missingness

Assumptions

- Denote $s_c = (\tau_1 = \dots = \tau_K = 1)$, missing pattern for “completers”
- Let Y_{obs} and Y_{mis} denote the observed and missing outcomes
- Benchmark assumptions
 - **CCMV**: Complete case missing-variable restrictions
 - For all s ,
$$f(Y_{\text{mis}}|Y_{\text{obs}}, X, T, S = s) = f(Y_{\text{mis}}|Y_{\text{obs}}, X, T, S = s_c)$$

Step 1: Model completers

- Denote (Y_1, \dots, Y_k) by \bar{Y}_k
- Factorize the joint distribution of \bar{Y}_K as

$$\begin{aligned} f(\bar{Y}_K | Y_0, X, T, S = s_c) \\ = \prod_{k=1}^K f(Y_k | \bar{Y}_{k-1}, Y_0, X, T, S = s_c) \end{aligned}$$

- Specify

$$\begin{aligned} Y_k | \bar{Y}_{k-1}, Y_0, X, T, S = s_c \\ = \alpha_{0,k}^T + \alpha_{1,k}^T \bar{Y}_{k-1} + \alpha_{2,k}^T Y_0 + \alpha_{3,k}^T X + \epsilon. \end{aligned}$$

- Allow ϵ to be **non-parametrically** distributed

Step 2: Impute missing data

- Under **normality** assumptions, $f(Y_{\text{mis}}|Y_{\text{obs}}, X, T, S = s_c)$ available in closed form
- Under **non-parametric** distribution assumptions, $f(Y_{\text{mis}}|Y_{\text{obs}}, X, T, S = s_c)$ can be numerically evaluated

Example: $K = 2$

S	τ_1	τ_2
s_1	0	0
s_2	0	1
s_3	1	0
s_4	1	1

Models for completers

$$f(Y_1|Y_0, X, T, S = s_4)$$

$$f(Y_2|Y_1, Y_0, X, T, S = s_4)$$

Imputation

$$f(Y_2, Y_1|Y_0, X, T, S = s_1) = f(Y_2, Y_1|Y_0, X, T, S = s_4)$$

$$f(Y_1|Y_2, Y_0, X, T, S = s_2) = f(Y_1|Y_2, Y_0, X, T, S = s_4)$$

$$f(Y_2|Y_1, Y_0, X, T, S = s_3) = f(Y_2|Y_1, Y_0, X, T, S = s_4)$$

Example: $K = 2$

- Consider a subject with Y_0, X, T , and $S = s_2$ (only Y_2 observed)
- Need to impute Y_1 from $f(Y_1|Y_2, Y_0, X, T, S = s_2)$
- By CCMV,
 $f(Y_1|Y_2, Y_0, X, T, S = s_2) = f(Y_1|Y_2, Y_0, X, T, S = s_4)$
- $f(Y_1|Y_2, Y_0, X, T, S = s_4)$ not available under **non-parametric** error distribution assumption

Example: $K = 2$

To sample from $f(Y_1|Y_2)$ (omit the condition on $Y_0, X, T, S = s_4$ for compactness), note

$$f(Y_1|Y_2) \propto f(Y_2|Y_1)f(Y_1)$$

- $f(Y_2|Y_1)$ bounded by M obtained by kernel density estimation
- rejection sampling using $f(Y_1)$ as an instrumental distribution

Sensitivity analysis

- Introduce sensitivity parameters Δ in a parsimonious way
- Alternative assumptions: for all s ,

$$f(Y_{\text{mis}}|Y_{\text{obs}}, X, T, S = s) \\ \propto \exp\{\Delta Z\} f(Y_{\text{mis}}|Y_{\text{obs}}, X, T, S = s_c)$$

Example

- $K = 2$
- $Z = (Y_2 + Y_1)/2 - Y_0$
- Given Δ , imputation assumption:

$$f(Y_2, Y_1 | Y_0, X, T, S = s_1) \propto e^{\Delta \frac{Y_2}{2}} e^{\Delta \frac{Y_1}{2}} f(Y_2, Y_1 | Y_0, X, T, S = s_4)$$

$$f(Y_1 | Y_2, Y_0, X, T, S = s_2) \propto e^{\Delta \frac{Y_1}{2}} f(Y_1 | Y_2, Y_0, X, T, S = s_4)$$

$$f(Y_2 | Y_1, Y_0, X, T, S = s_3) \propto e^{\Delta \frac{Y_2}{2}} f(Y_2 | Y_1, Y_0, X, T, S = s_4)$$

Exponential tilting model

Consider an exponential tilting model

$$f_{Y'}(y) \propto e^{\Delta y} f_Y(y)$$

- Under normality
 - $Y \sim N(\mu, \sigma^2)$
 - $Y' \sim N(\mu + \Delta\sigma^2, \sigma^2)$
- Under non-parametric assumption
 - $\hat{f}_Y(y) = \sum_{i=1}^n \frac{1}{n} K_h(y - Y_i)$
 - $\hat{f}_{Y'}(y) = \sum_{i=1}^n \frac{e^{\Delta Y_i}}{\sum_{j=1}^n e^{\Delta Y_j}} K_h(y - Y_i)$

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Web application

- Currently available at
<http://ebayes.synology.me/shiny/composite/>
- Major components
 - upload and review data
 - specify endpoints and imputation model
 - basic graphics
 - specify ranking rule
 - generate imputed dataset
 - bootstrap analysis

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Analysis

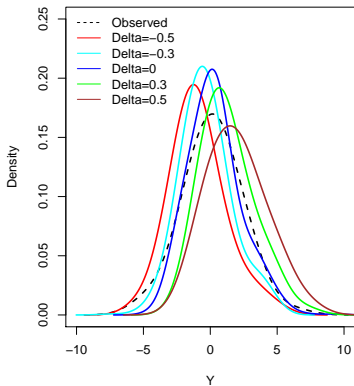
- Covariates

Covariates	Levels
ECOG	0:{0, 1}, 1:{2}
AGE	0:≤ 65, 1:> 65
SEX	0:M, 1:F
BMI	0:≤ 18.5, 1:> 18.5
WEIGHT LOSS	0:≤ 10%, 1:> 10%
Y0	Continuous

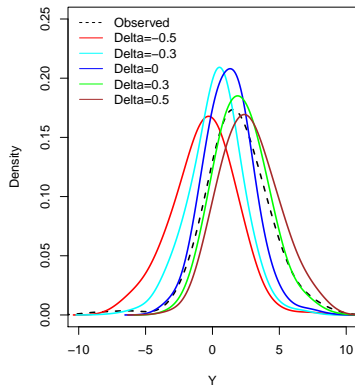
- 500 bootstrap samples, 15 imputed datasets for each bootstrap sample
- Sensitivity parameters $\Delta = \{-0.5, -0.4, \dots, 0.5\}$

Imputed data

Arm 0



Arm 1



Hypothesis testing

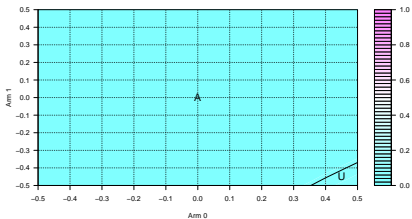
Normality	$\hat{\theta}(95\%CI)$	p-value
<i>Without Normality</i>	0.28(0.17, 0.38)	< 0.0001
<i>With Normality</i>	0.24(0.13, 0.35)	< 0.0001

Median

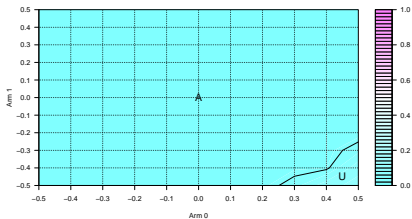
Normality	Arm 0 (95%CI)	Arm 1 (95%CI)
<i>Without Normality</i>	-0.44(-0.88, 0.20)	1.10(0.76, 1.42)
<i>With Normality</i>	-0.49(-1.09, 0.22)	1.03(0.62, 1.36)

Sensitivity analysis

Without normality



With normality



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Summary

- Propose a composite endpoint approach for evaluating treatment effects in randomized clinical trials with **death** and **missingness**
- Apply complete case missing-variable restrictions (**CCMV**) for handling missing data in survivors
- Exponential tilting model for sensitivity analysis
- Online web application developed

THE END