Global Sensitivity Analysis of Randomized Trials with Missing Data ICTR Lunch and Learn

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- FDA
- PCORI
- NIH Project DIDACT

Missing Data Matters

- While unbiased estimates of treatment effects can be obtained from randomized trials with no missing data, this is no longer true when data are missing on some patients.
- The essential problem is that inference about treatment effects relies on *unverifiable* assumptions about the nature of the mechanism that generates the missing data.
- While we usually know the reasons for missing data, we do not know the distribution of outcomes for patients with missing data, how it compares to that of patients with observed data and whether differences in these distributions can be explained by the observed data.

Robert Temple and Bob O'Neil (FDA)

 "During almost 30 years of review experience, the issue of missing data in ... clinical trials has been a major concern because of the potential impact on the inferences that can be drawn when data are missing the analysis and interpretation of the study pose a challenge and the conclusions become more tenuous as the extent of 'missingness' increases."

NRC Report and Sensitivity Analysis

- In 2010, the National Research Council (NRC) issued a reported entitled "The Prevention and Treatment of Missing Data in Clinical Trials."
- This report, commissioned by the FDA, provides 18 recommendations targeted at (1) trial design and conduct, (2) analysis and (3) directions for future research.
- Recommendation 15 states
 - Sensitivity analyses should be part of the primary reporting of findings from clinical trials. Examining sensitivity to the assumptions about the missing data mechanism should be a mandatory component of reporting.

ICH, EMEA and Sensitivity Analysis

- 1998 International Conference of Harmonization (ICH) Guidance document (E9) entitled "Statistical Principles in Clinical Trials" states: "*it is important to evaluate the robustness of the results to various limitations of the data, assumptions, and analytic approaches to data analysis*"
- European Medicines Agency 2009 draft "Guideline on Missing Data in Confirmatory Clinical Trials" states "[i]n all submissions with non-negligible amounts of missing data sensitivity analyses should be presented as support to the main analysis."

PCORI and Sensitivity Analysis

- In 2012, Li *et al.* issued the report "Minimal Standards in the Prevention and Handling of Missing Data in Observational and Experimental Patient Centered Outcomes Research"
- This report, commissioned by PCORI, provides 10 standards targeted at (1) design, (2) conduct, (3) analysis and (4) reporting.
- Standard 8 echoes the NRC report, stating
 - Examining sensitivity to the assumptions about the missing data mechanism (i.e., sensitivity analysis) should be a mandatory component of the study protocol, analysis, and reporting.

The set of possible assumptions about the missing data mechanism is very large and cannot be fully explored. There are different approaches to sensitivity analysis:

- Ad-hoc
- Local
- Global

Ad-hoc Sensitivity Analysis

- Analyzing data using a few different analytic methods, such as last or baseline observation carried forward, complete or available-case analysis, mixed models or multiple imputation, and evaluate whether the resulting inferences are consistent.
- The problem with this approach is that the assumptions that underlie these methods are very strong and for many of these methods unreasonable.
- More importantly, just because the inferences are consistent does not mean that there are no other reasonable assumptions under which the inference about the treatment effect is different.

- Specify a reasonable benchmark assumption (e.g., missing at random) and evaluate the robustness of the results within a small neighborhood of this assumption.
- What if there are assumptions outside the local neighborhood which are plausible?

- Evaluate robustness of results across a much broader range of assumptions that include a reasonable benchmark assumption and a collection of additional assumptions that trend toward best and worst case assumptions.
- Emphasized in Chapter 5 of the NRC report.
- This approach is substantially more informative because it operates like "stress testing" in reliability engineering, where a product is systematically subjected to increasingly exaggerated forces/conditions in order to determine its breaking point.

- In the missing data setting, global sensitivity analysis allows one to see how far one needs to deviate from the benchmark assumption in order for inferences to change.
- "Tipping point" analysis (Yan, Lee and Li, 2009; Campbell, Pennello and Yue, 2011)
- If the assumptions under which the inferences change are judged to be sufficiently far from the benchmark assumption, then greater credibility is lent to the benchmark analysis; if not, the benchmark analysis can be considered to be fragile.

- Restrict consideration to follow-up randomized study designs that prescribe that measurements of an outcome of interest are to be taken on each study participant at fixed time-points.
- Focus on monotone missing data pattern
- Consider the case where interest is focused on a comparison of treatment arm means at the last scheduled visit.

Case Study: Quetiapine Bipolar Trial

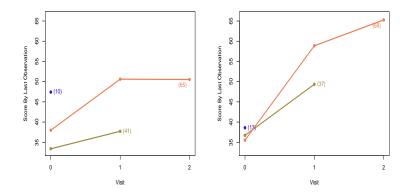
- Patients with bipolar disorder randomized equally to one of three treatment arms: placebo, Quetiapine 300 mg/day or Quetiapine 600 mg/day (Calabrese *et al.*, 2005).
- Randomization was stratified by type of bipolar disorder.
- Short-form version of the Quality of Life Enjoyment Satisfaction Questionnaire (QLESSF, Endicott *et al.*, 1993), was scheduled to be measured at baseline, week 4 and week 8.
- Focus on the subset of 234 patients with bipolar 1 disorder who were randomized to either the placebo (n=116) or 600 mg/day (n=118) arms.

Quetiapine Bipolar Trial

- 600 mg/day dose was titrated to achieve target by Day 8.
- In each treatment group, a dose reduction of 100 mg was allowed to improve tolerability.
- At discretion of the investigator, patients could be discontinued from study treatment and assessments at any time.
- Patients were free to discontinue their participation in the study at any time.
- Use of psychoactive drugs, except lorazepam and zolpidem tartrate during the first 3 weeks, was prohibited. Investigators were allowed to prescribe other medications for the safety and well-being of the participant.

- Only 65 patients (56%) in placebo arm and 68 patients (58%) in the 600mg/day arm had a complete set of QLESSF scores.
- Patients with complete data tend to have higher average QLESSF scores, suggesting that a complete-case analysis could be biased.

Figure: Treatment-specific (left: placebo; right: 600 mg/day Quetiapine) trajectories of mean QLESSF scores, stratified by last available measurement.



What is the difference in the mean QLESSF score at week 8 between Quetiapine 600 mg/day and placebo in the counterfactual world in which all patients were followed to that week?

Imagination

- Validity of assumptions will depend on what is *imagined* about treatments that patients receive off-study.
- *Not* imagining the continuation of assigned treatment after occurrence of intolerable side effects or lack of efficacy.
- Imagining that patients receive treatment as close to the assigned treatment as ethically possible.
- The difference of the treatment-specific mean QLESSF outcomes at week 8 under this imaginary, yet plausible, treatment scenario is the target *estimand* of interest.

- Inference about the treatment arm means requires two types of assumptions:
 - (i) *unverifiable* assumptions about the distribution of outcomes among those with missing data and
 - (ii) additional testable assumptions that serve to increase the efficiency of estimation.

- Type (i) assumptions are necessary to identify the treatment-specific means.
- By *identification*, we mean that we can write it as a function that depends only on the distribution of the observed data.
- When a parameter is identified we can hope to estimate it as precisely as we desire with a sufficiently large sample size,
- In the absence of identification, statistical inference is fruitless as we would be unable to learn about the true parameter value even if the sample size were infinite.

- To address the identifiability issue, it is essential to conduct a sensitivity analysis, whereby the data analysis is repeated under different type (i) assumptions, so as to investigate the extent to which the conclusions of the trial are dependent on these subjective, unverifiable assumptions.
- The usefulness of a sensitivity analysis ultimately depends on the plausibility of the unverifiable assumptions.
- It is key that any sensitivity analysis methodology allow the formulation of these assumptions in a transparent and easy to communicate manner.

- There are an infinite number of ways of positing type (i) assumptions.
- Ultimately, however, these assumptions prescribe how missing outcomes should be "imputed."
- A reasonable way to posit these assumptions is to
 - stratify individuals with missing outcomes according to the data that we were able to collect on them and the occasions at which the data were collected
 - separately for each stratum, hypothesize a connection (or link) between the distribution of the missing outcome with the distribution of the outcome among those with the observed outcome and who share the same recorded data.

- Type (i) assumptions will not suffice when the repeated outcomes are continuous or categorical with many levels. This is because of *data sparsity*.
- For example, the stratum of people who share the same recorded data will typically be small. As a result, it is necessary to draw strength across strata by "smoothing."
- Without smoothing, the data analysis will rarely be informative because the uncertainty concerning the treatment arm means will often be too large to be of substantive use.
- As a result, it is necessary to impose type (ii) smoothing assumptions.
- Type (ii) assumptions should be scrutinized with standard model checking techniques.

- The global sensitivity framework proceeds by parameterizing (i.e., indexing) the connections (i.e., type (i) assumptions) via sensitivity analysis parameters.
- The parameterization is configured so that a specific value of the sensitivity analysis parameters (typically set to zero) corresponds to a benchmark connection that is considered reasonably plausible and sensitivity analysis parameters further from the benchmark value represent more extreme departures from the benchmark connection.
- The global sensitivity analysis strategy that we propose is focused on separate inferences for each treatment arm, which are then combined to evaluate treatment effects.

Notation: Quetiapine Bipolar Trial

- Y₀, Y₁, Y₂: QLESSF scores scheduled to be collected at baseline, week 4 and week 8.
- Let R_k be the indicator that Y_k is observed.
- We assume R₀ = 1 and that R_k = 0 implies R_{k+1} = 0 (i.e., missingness is monotone).
- Patient is on-study at visit k if $R_k = 1$
- Patient discontinued prior to visit k if $R_k = 0$
- Patient last seen at visit k 1 if $R_{k-1} = 1$ and $R_k = 0$.
- Y_k^{obs} equals to Y_k if $R_k = 1$ and equals to *nil* if $R_k = 0$.

Notation: Quetiapine Bipolar Trial

• The observed data for an individual are

$$O = (Y_0, R_1, Y_1^{obs}, R_2, Y_2^{obs}),$$

- The distribution of the observed data can be represented by:
 - *dist*(*Y*₀)
 - $prob(R_1 = 1|Y_0)$
 - $dist(Y_1|R_1 = 1, Y_0)$
 - $prob(R_2 = 1 | R_1 = 1, Y_1, Y_0)$
 - $dist(Y_2|R_2 = 1, Y_1, Y_0).$
- We assume that we observed *n* copies of *O*.
- The goal is to use these data to draw inference about the mean of *Y*₂.

- $A_0(y_0)$: patients last seen at visit 0 ($R_0 = 1, R_1 = 0$) with $Y_0 = y_0$.
- $B_1(y_0)$: patients on-study at visit 1 ($R_1 = 1$) with $Y_0 = y_0$.
- $A_1(y_0, y_1)$: patients last seen at visit 1 ($R_1 = 1, R_2 = 0$) with $Y_0 = y_0$ and $Y_1 = y_1$.
- $B_2(y_0, y_1)$: patients who complete study $(R_2 = 1)$ with $Y_0 = y_0 \ Y_1 = y_1$.

Missing at random posits the following type (i) "linking" assumptions:

- For each y₀, the distribution of Y₁ and Y₂ is the same for those in stratum A₀(y₀) as those in stratum B₁(y₀).
- For each y₀, y₁, the distribution of Y₂ is the same for those in stratum A₁(y₀, y₁) as those in stratum B₂(y₀, y₁).

Mathematically, we can express these assumptions as follows:

$$dist(Y_1, Y_2|A_0(y_0)) = dist(Y_1, Y_2|B_1(y_0))$$
 for all y_0 (1)

and

$$dist(Y_2|A_1(y_0, y_1)) = dist(Y_2|B_2(y_0, y_1))$$
 for all y_0, y_1 (2)

$$prob(R_1 = 0 | R_0 = 1, Y_0 = y_0, Y_1 = y_1, Y_2 = y_2)$$

= $prob(R_1 = 0 | R_0 = 1, Y_0 = y_0)$

and

$$prob(R_2 = 0 | R_1 = 1, Y_0 = y_0, Y_1 = y_1, Y_2 = y_2)$$

= $prob(R_2 = 0 | R_1 = 1, Y_0 = y_0, Y_1 = y_1)$

Missing at random implies:

- The decision to discontinue the study before visit 1 is like the flip of a coin with probability depending on the value of the outcome at visit 0.
- For those on-study at visit 1, the decision to discontinue the study before visit 2 is like the flip of a coin with probability depending on the value of the outcomes at visits 1 and 0.

- MAR is a type (i) assumption. It is "unverifiable."
- For patients last seen at visit k, we cannot learn from the observed data about the conditional (on observed history) distribution of outcomes after visit k.
- For patients last seen at visit k, any assumption that we make about the conditional (on observed history) distribution of the outcomes after visit k will be unverifiable from the data available to us.
- For patients last seen at visit k, the assumption that the conditional (on observed history) distribution of outcomes after visit k is the same as those who remain on-study after visit k is unverifiable.

- Under MAR, the mean of Y_2 is identified.
- That is, it can be expressed as a function of the distribution of the observed data.

The MAR assumption is not the only one that is (a) unverifiable and (b) allows identification of the mean of Y_2 .

Missing Not at Random (MNAR)

The first part of the MAR assumption (see (1) above) is

 $dist(Y_1, Y_2|A_0(y_0)) = dist(Y_1, Y_2|B_1(y_0))$ for all y_0

It is equivalent to

$$dist(Y_2|A_0(y_0), Y_1 = y_1) = dist(Y_2|B_1(y_0), Y_1 = y_1) \text{ for all } y_0, y_1$$
(3)

and

$$dist(Y_1|A_0(y_0)) = dist(Y_1|B_1(y_0))$$
 for all y_0 (4)

In building a class of MNAR models, we will retain (3):

- For all y₀, y₁, the distribution of Y₂ for patients in stratum A₀(y₀) with Y₁ = y₁ is the same as the distribution of Y₂ for patients in stratum B₁(y₀) with Y₁ = y₁.
- The decision to discontinue the study before visit 1 is independent of Y_2 (i.e., the future outcome) after conditioning on the Y_0 (i.e., the past outcome) and Y_1 (i.e., the most recent outcome).
- Non-future dependence (Diggle and Kenward, 1994)

Missing Not at Random (MNAR)

Generalize (4) using exponential tilting

 $dist(Y_1|A_0(y_0))$ $\propto dist(Y_1|B_1(y_0)) \exp\{\alpha r(Y_1)\} \text{ for all } y_0$ (5)

Generalize (2) using exponential tilting

 $dist(Y_2|A_1(y_0, y_1)) \\ \propto dist(Y_2|B_2(y_0, y_1)) \exp\{\alpha r(Y_2)\} \text{ for all } y_0, y_1 \quad (6)$

- r(y) is a specified increasing function; α is a sensitivity analysis parameter.
- $\alpha = 0$ is MAR.

When $\alpha > 0$ (< 0)

- For each y₀, the distribution of Y₁ for patients in stratum A₀(y₀) is weighted more heavily to higher (lower) values than the distribution of Y₁ for patients in stratum B₁(y₀).
- For each y₀, y₁, the distribution of Y₂ for patients in stratum A₁(y₀, y₁) is weighted more heavily to higher (lower) values than the distribution of Y₂ for patients in stratum B₂(y₀, y₁).

The amount of "tilting" increases with the magnitude of α .

Missing Not at Random (MNAR)

logit
$$prob(R_1 = 0 | R_0 = 1, Y_0 = y_0, Y_1 = y_1, Y_2 = y_2)$$

= $l_1(y_0) + \alpha r(y_1)$

and

logit
$$prob(R_2 = 0 | R_1 = 1, Y_0 = y_0, Y_1 = y_1, Y_2 = y_2)$$

= $l_2(y_0, y_1) + \alpha r(y_2)$

Written in this way:

- The decision to discontinue the study before visit 1 is like the flip of a coin with probability depending on the value of the outcome at visit 0 *and (in a specified way)* the value of the outcome at visit 1.
- For those on-study at visit 1, the decision to discontinue the study before visit 2 is like the flip of a coin with probability depending on the value of the outcomes at visits 0 and 1 and (in a specified way) the value of the outcome at visit 2.

Exponential Tilting Explained

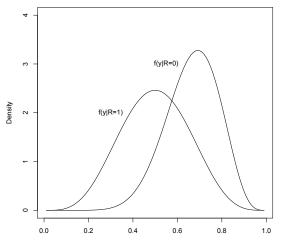
$$dist(Y|R=0) \propto dist(Y|R=1) \exp\{\alpha r(Y)\}$$

• If
$$dist(Y|R = 1) \sim N(\mu, \sigma^2)$$
 and $r(Y) = Y$,
 $dist(Y|R = 0) \sim N(\mu + \alpha \sigma^2, \sigma^2)$

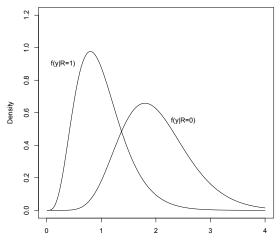
- If $dist(Y|R = 1) \sim Beta(a, b)$ and r(Y) = log(Y), $dist(Y|R = 0) \sim Beta(a + \alpha, b)$, $\alpha > -a$.
- If $dist(Y|R = 1) \sim Gamma(a, b)$ and $r(Y) = \log(Y)$, $dist(Y|R = 0) \sim Gamma(a + \alpha, b)$, $\alpha > -a$.
- If $dist(Y|R = 1) \sim Gamma(a, b)$ and r(Y) = Y, $dist(Y|R = 0) \sim Gamma(a, b - \alpha)$, $\alpha < b$.

• If
$$dist(Y|R = 1) \sim Bernoulli(p)$$
 and $r(Y) = Y$,
 $dist(Y|R = 0) \sim Bernoulli\left(\frac{p \exp(\alpha)}{p \exp(\alpha)+1-p}\right)$.

Beta



Gamma

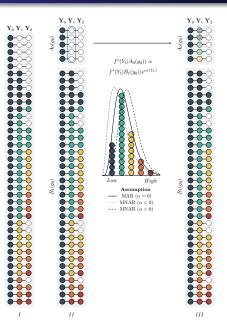


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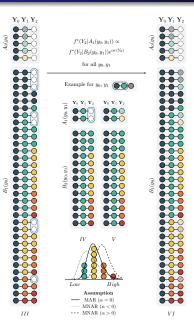
Missing Not at Random (MNAR)

- For given α , the mean of Y_2 is identified.
- That is, it can be expressed as a function of the distribution of the observed data.

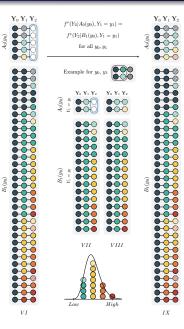
Global Sensitivity Analysis



Global Sensitivity Analysis



Global Sensitivity Analysis



Inference

For given α , the identification formula depends on

- $dist(Y_2|R_2 = 1, Y_1 = y_1, Y_0 = y_0)$
- $dist(Y_1|R_1 = 1, Y_0 = y_0)$
- dist(Y₀)
- $prob(R_2 = 0 | R_1 = 1, Y_1 = y_1, Y_0 = y_0)$

•
$$prob(R_1 = 0 | R_0 = 1, Y_0 = y_0).$$

It is natural to consider estimating the mean of Y_2 by "plugging in" estimators of these quantities.

How can we estimate these latter quantities? With the exception of $dist(Y_0)$, it is tempting to think that we can use non-parametric procedures to estimate these quantities.

- Plug-in estimator will perform poorly.
- Need to use
 - type (ii) smoothing assumptions
 - different estimation strategy (plug-in correction procedure)

Inference - Type (ii) Assumptions

$$dist(Y_2|R_2 = 1, Y_1 = y_1, Y_0 = y_0) = dist(Y_2|R_2 = 1, Y_1 = y_1)$$
(7)

and

$$prob(R_2 = 0 | R_1 = 1, Y_1 = y_1, Y_0 = y_0) = prob(R_2 = 0 | R_1 = 1, Y_1 = y_1)$$
(8)

Estimate

- $dist(Y_2|R_2 = 1, Y_1 = y_1)$
- $dist(Y_1|R_1 = 1, Y_0 = y_0)$
- $prob(R_2 = 0 | R_1 = 1, Y_1 = y_1)$
- $prob(R_1 = 0 | R_0 = 1, Y_0 = y_0).$

using kernel smoothing technique (with cross validation).

Quetiapine Bipolar Trial - Fit

- Estimated smoothing parameters for the drop-out model are 11.54 and 9.82 for the placebo and 600 mg arms.
- Estimated smoothing parameters for the outcome model are 6.34 and 8.05 for the placebo and 600 mg arms.
- In the placebo arm, the observed percentages of last being seen at visits 0 and 1 among those at risk at these visits are 8.62% and 38.68%. Model-based estimates are 7.99% and 38.19%.
- For the 600 mg arm, the observed percentages are 11.02% and 35.24% and the model-based estimates are 11.70% and 35.08%.

- In the placebo arm, the Kolmogorov-Smirnov distances between the empirical distribution of the observed outcomes and the model-based estimates of the distribution of outcomes among those on-study at visits 1 and 2 are 0.013 and 0.033.
- In the 600 mg arm, these distances are 0.013 and 0.022.
- These results suggest that our model for the observed data fits the observed data well.

- Under MAR, the estimated values of mean of Y₂ are 46.45 (95% CI: 42.35,50.54) and 62.87 (95% CI: 58.60,67.14) for the placebo and 600 mg arms.
- The estimated difference between 600 mg and placebo is 16.42 (95% 10.34, 22.51)
- Clinically significant improvement in quality of life in favor of Quetiapine.

- We set r(y) = y and ranged the sensitivity analysis parameter from -10 and 10 in each treatment arm.
- According to experts, there is no evidence to suggest that there is a differential effect of a unit change in QLESSF on the hazard of drop-out based on its location on the scale.

Figure: Treatment-specific (left: placebo; right: 600 mg/day Quetiapine) estimates (along with 95% pointwise confidence intervals) of μ^* as a function of α .

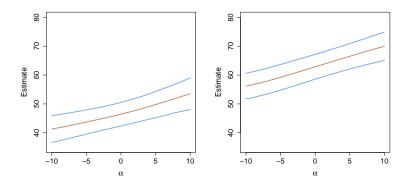


Figure: Contour plot of the estimated differences between mean QLESSF at Visit 2 for Quetiapine vs. placebo for various treatment-specific combinations of the sensitivity analysis parameters.

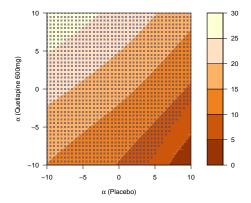
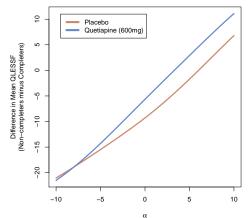
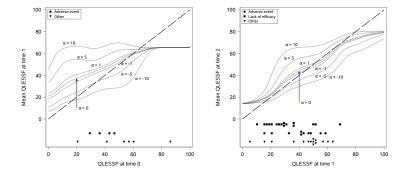
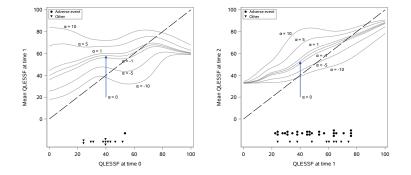


Figure: Treatment-specific differences between the estimated mean QLESSF at Visit 2 among non-completers and the estimated mean among completers, as a function of α .



• Only when the sensitivity analysis are highly differential (e.g., α (placebo) = 8 and α (Quetiapine) = -8) do 95% confidence intervals for treatment difference include zero.





• Conclusions under MAR are highly robust.

- Available on CRAN
- SAS and R versions available at

www.missingdatamatters.org

• Handles intermittent missing data

- No substitute for better trial design and procedures to minimize missing data.
- Global sensitivity analysis should be a mandatory component of trial reporting.
- Visit us at www.missingdatamatters.org or email me at dscharf@jhu.edu