

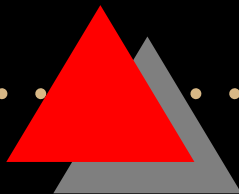


Bayesian Methods

LABORATORY

Lesson 1: Jan 24 2002

Software: R

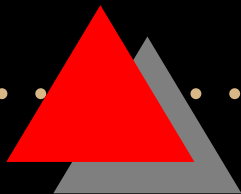




The R Project for Statistical Computing

<http://www.r-project.org/>

- R is a **language** and **environment** for statistical computing and graphics.

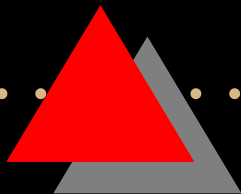




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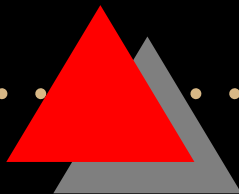
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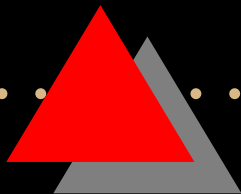
- R is a **language** and **environment** for statistical computing and graphics.
 - R, like S, is designed around a computer language, and it allows users to add additional functionality by defining new functions.
 - The term "environment" is intended to characterize it as a fully planned and coherent system, rather than an incremental accretion of very specific and inflexible tools.


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 - Free software is a matter of the users' freedom to run, copy, distribute, study, change and improve the software. It is not a matter of price!
 - R can be considered as **a different implementation of S**. There are some important differences, but much code written for S runs unaltered under R.



ONE-DIMENSIONAL *parameter models*

1. The Binomial model and its conjugate Beta prior

in $Bin(n, \theta)$ there is a single parameter of interest (n is typically assumed known), that is **the probability θ** of a certain outcome in each of the n trials considered.

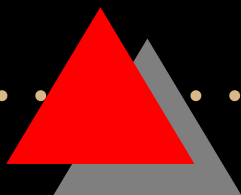


Bayesian estimation of a probability from BINOMIAL data

Gelman book, pag. 39, sec. 2.5

R code `placenta.r` is in the Lab notes at the course web page

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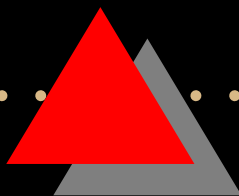


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- Our interest focus on the **proportion** of female births in the so called maternal condition *placenta previa*
- Our data consist in a early study in Germany: **437** females on **980** placenta previa births
- How much evidence do they provide that the proportion of placenta previa female births is < 0.485 , the proportion of the general population female births?



Analysis using a UNIFORM PRIOR

- Let the 1-parameter θ denote the proportion of *placenta previa* female births



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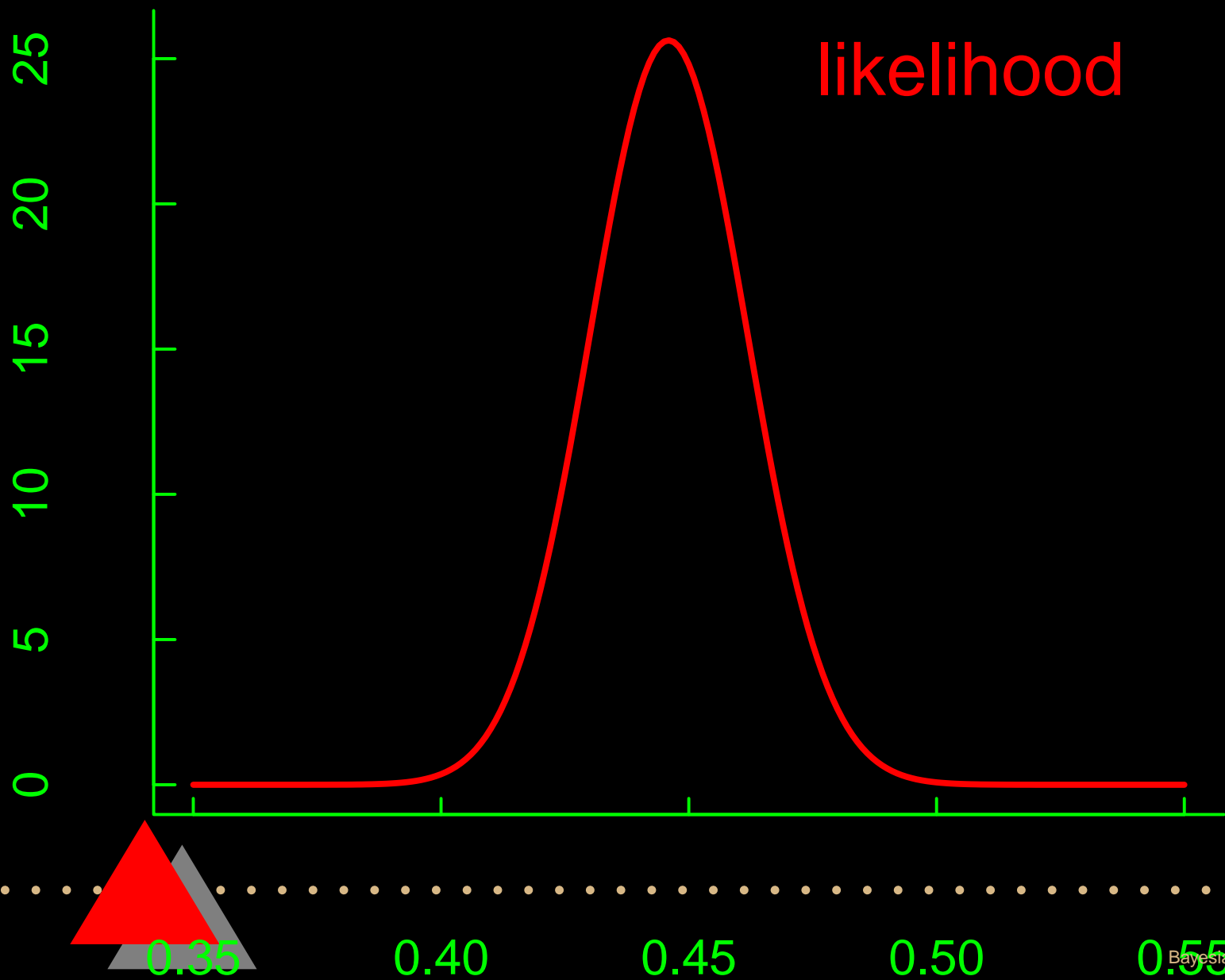
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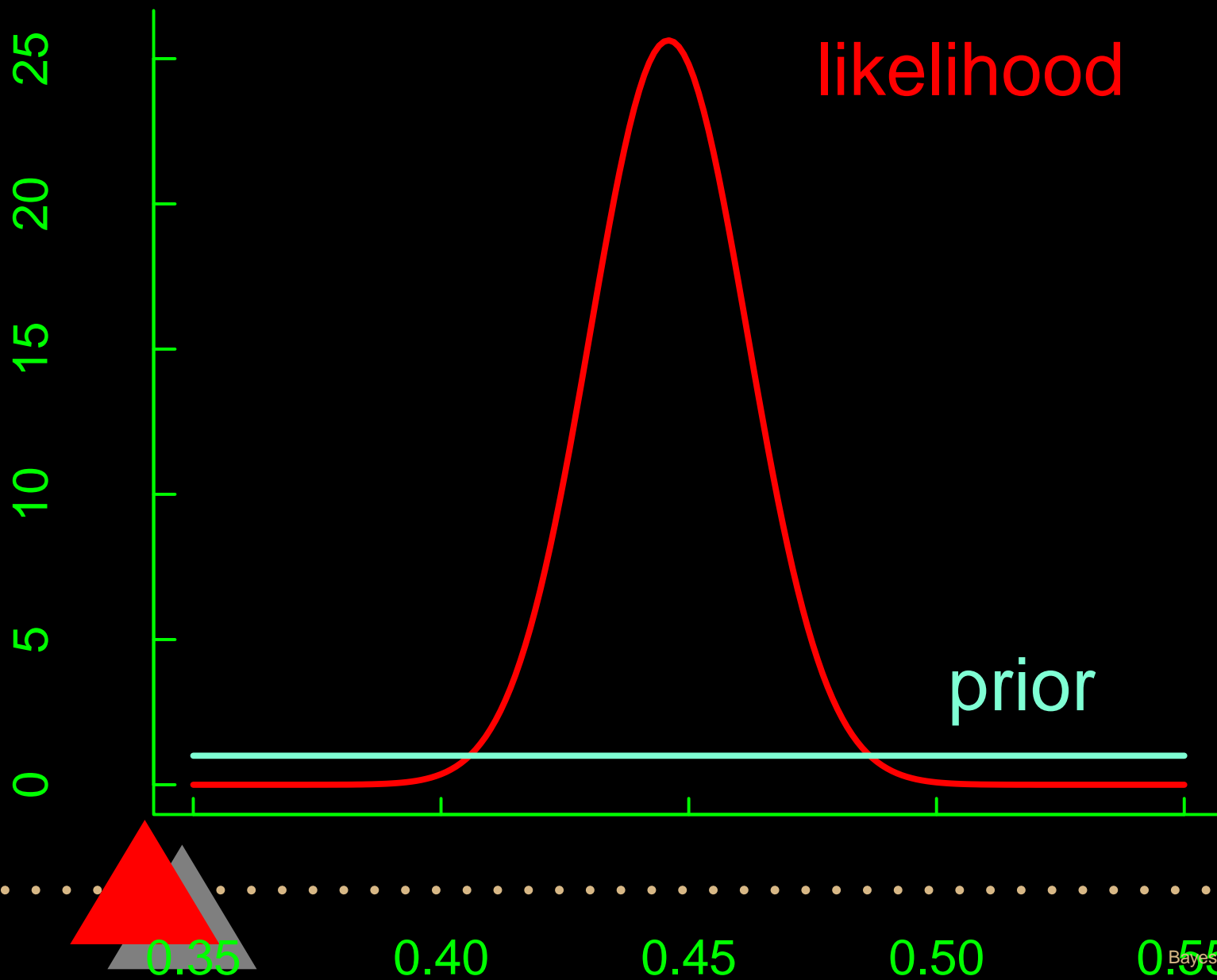
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- We specify the *prior* for θ to be a $U[0, 1]$
- The *posterior* for θ is, then, $\propto \theta^{437} (1 - \theta)^{980-437}$, i.e., is a $Beta(437 + 1, 980 - 437 + 1)$

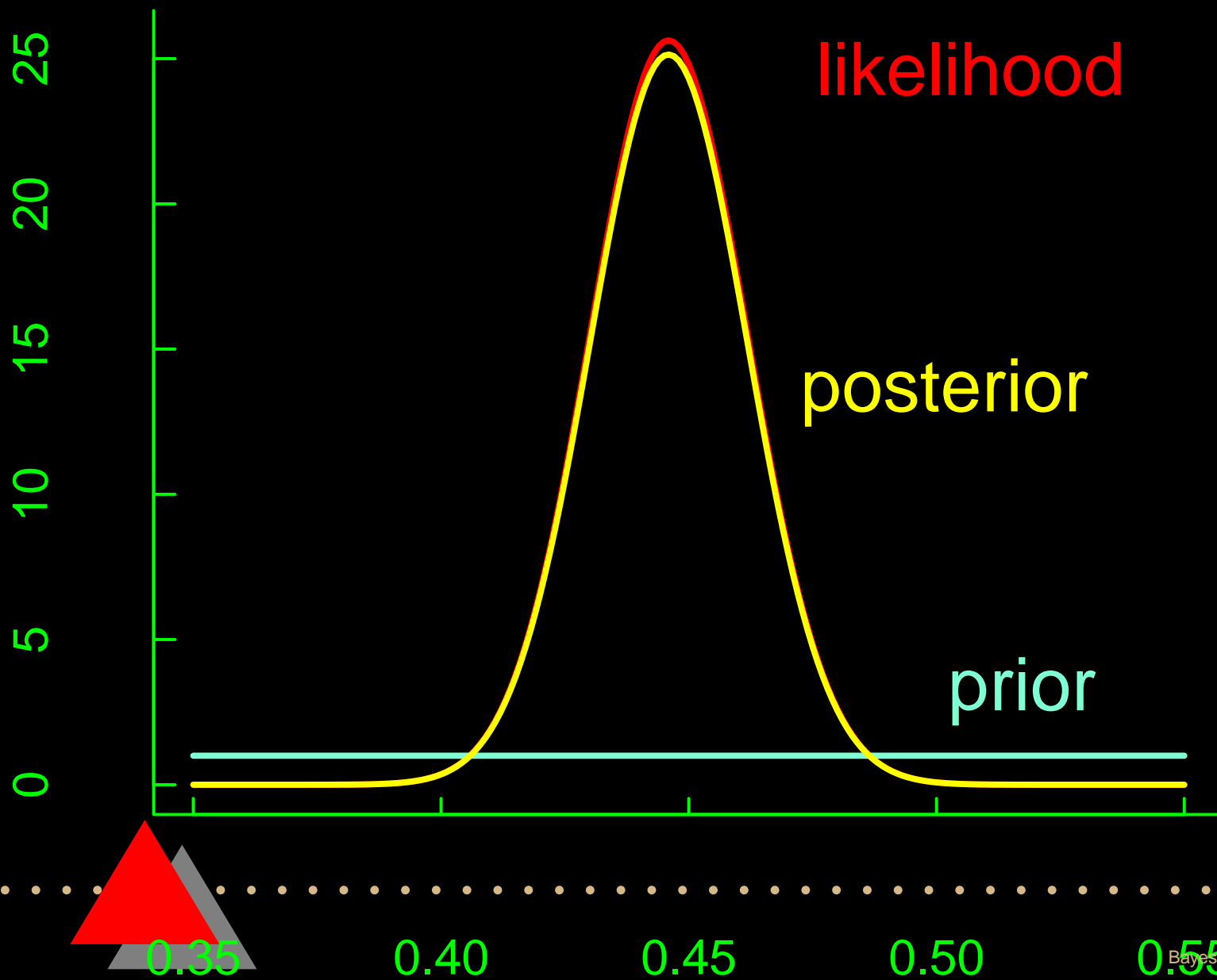
(Beta-)Uniform-Binomial



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Analysis using different BETA PRIORS

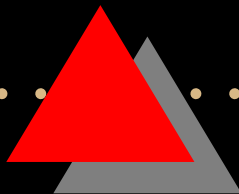
As the likelihood $p(y|\theta) \equiv L(\theta; y)$ is $\propto \theta^y (1 - \theta)^{n-y}$
if the prior is of the same form, e.g., $p(\theta)$ is \propto

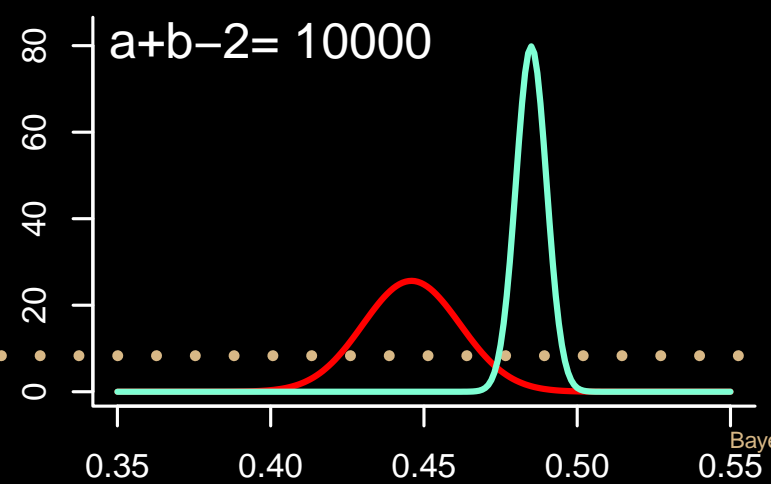
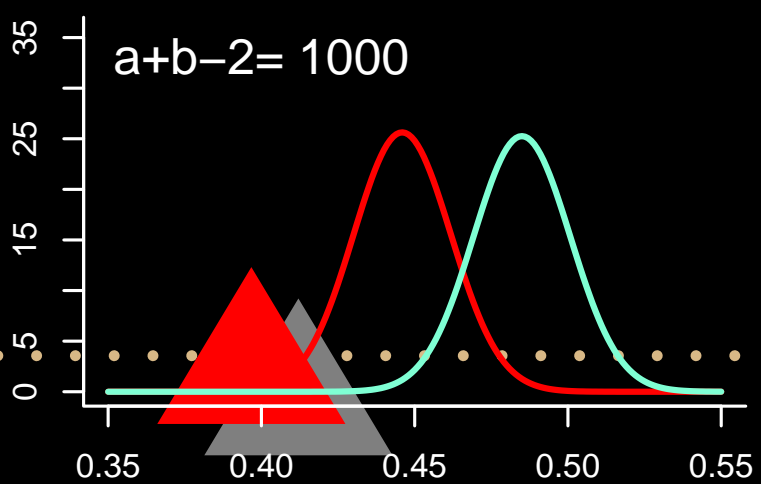
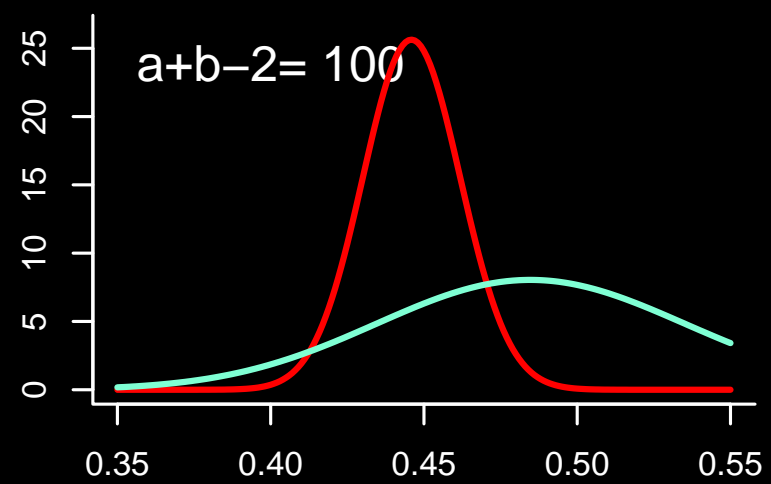
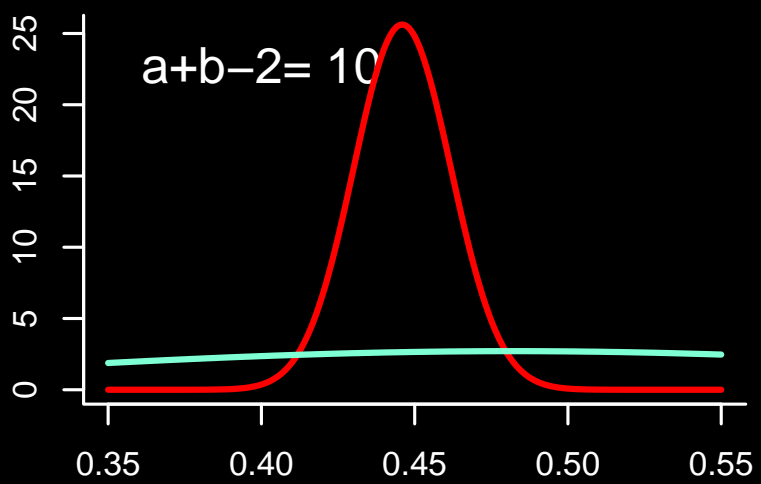
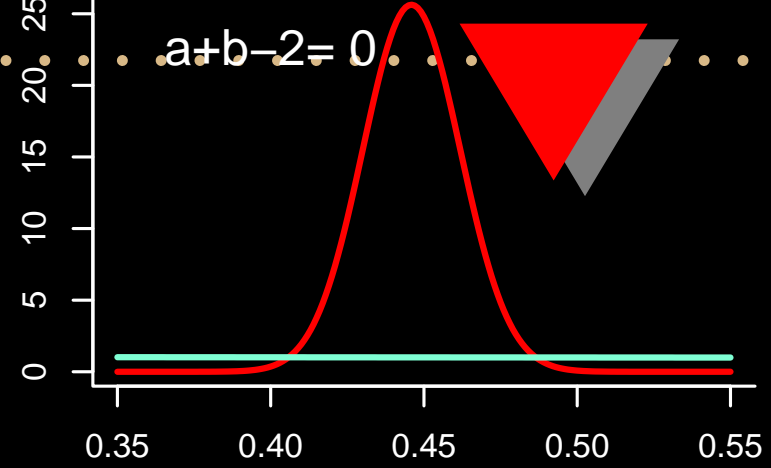
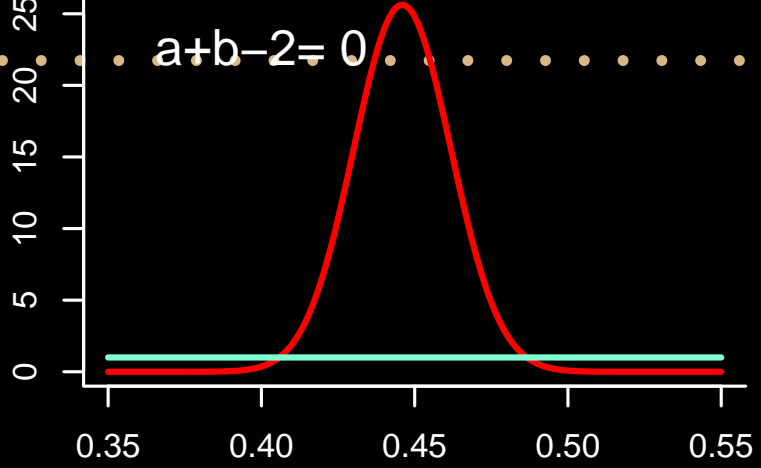
$$\theta^{\alpha-1} (1 - \theta)^{\beta-1}$$

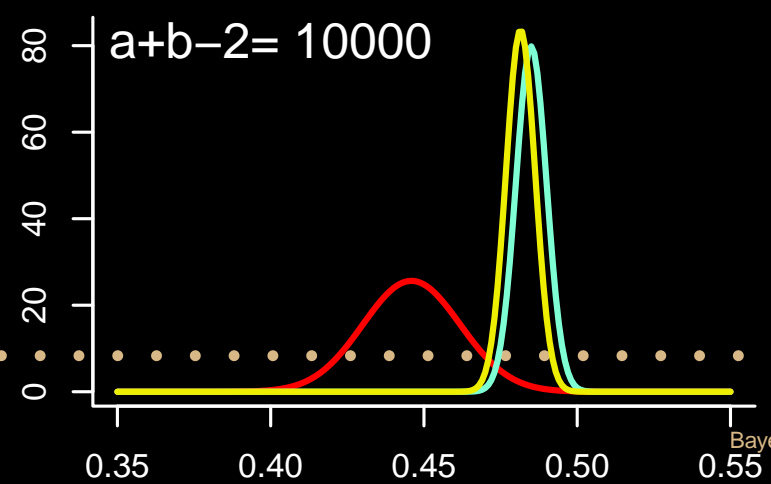
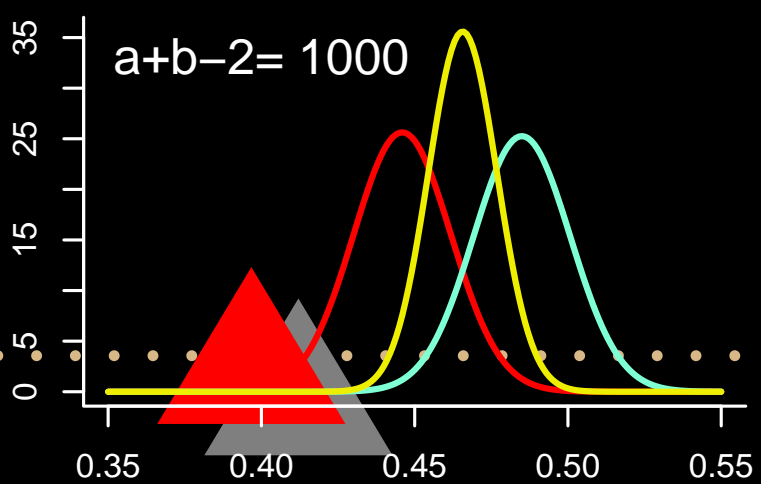
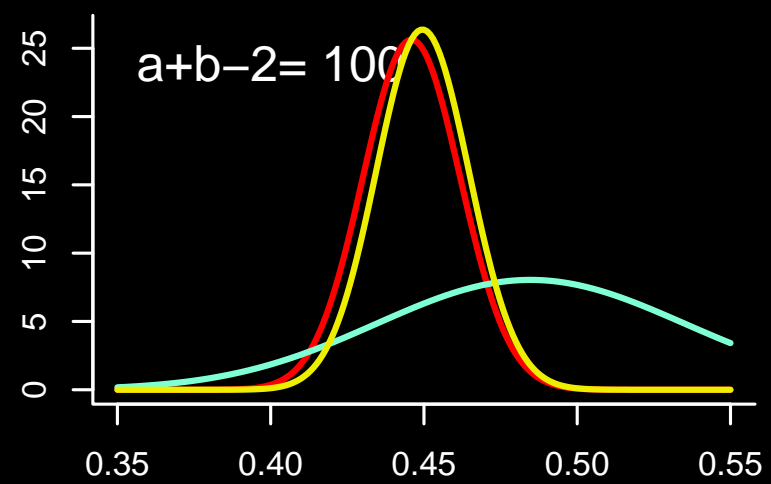
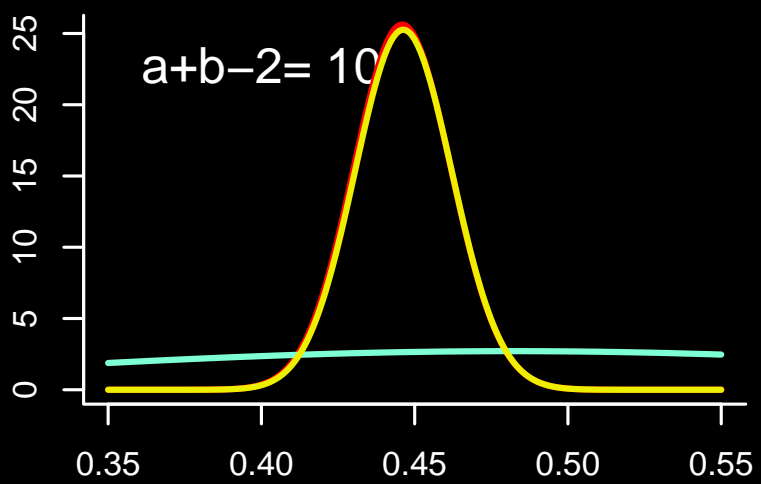
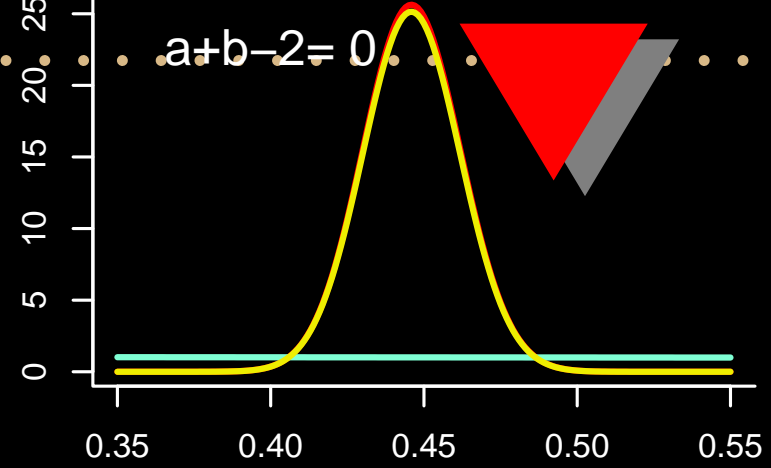
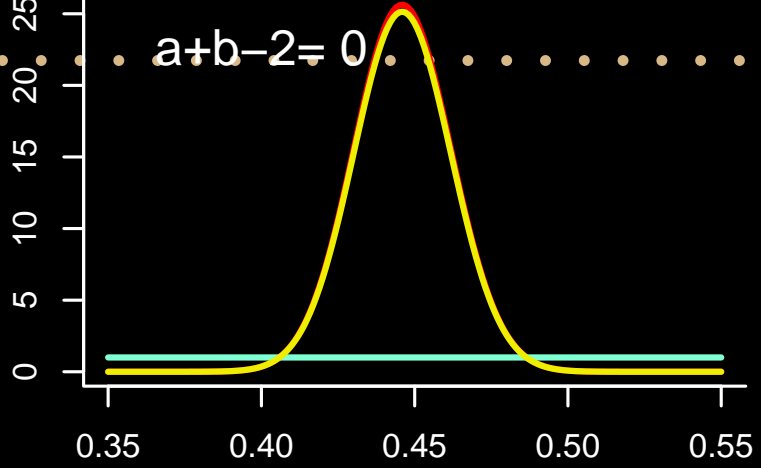
then the posterior will also be of this form. In fact, $p(\theta|y)$ is

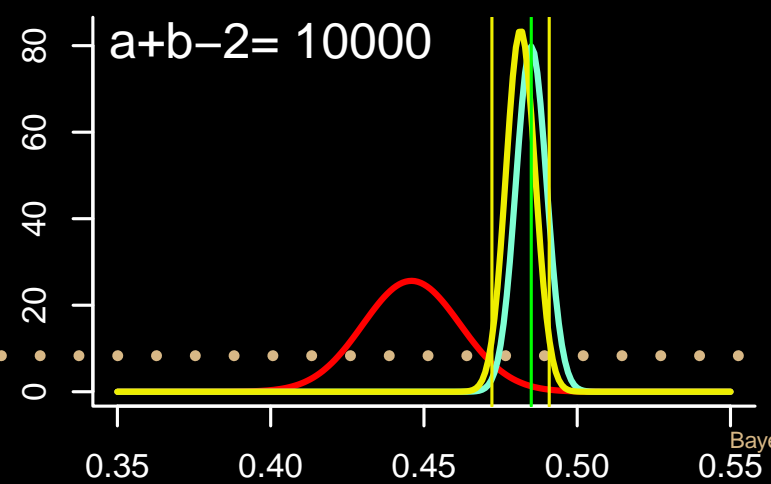
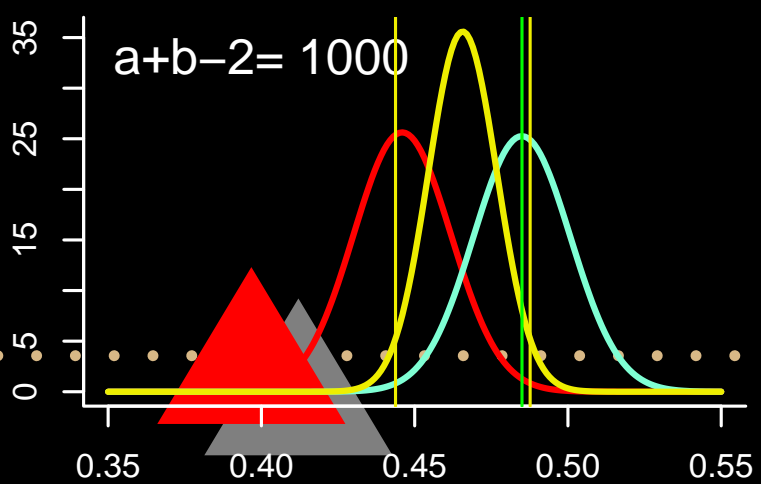
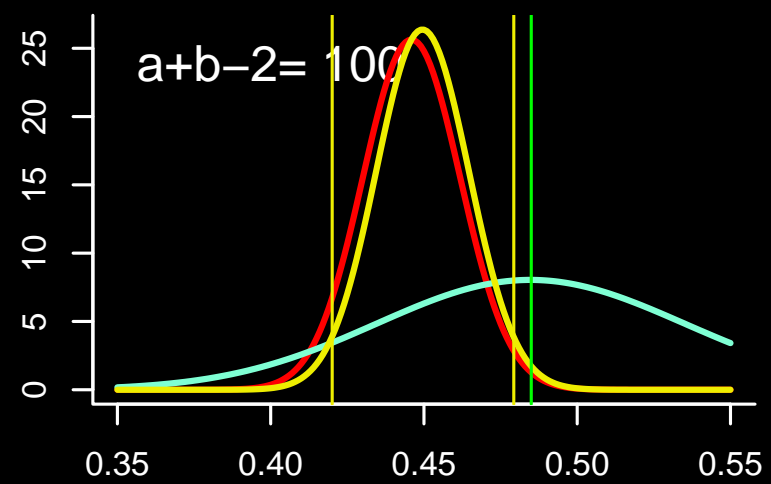
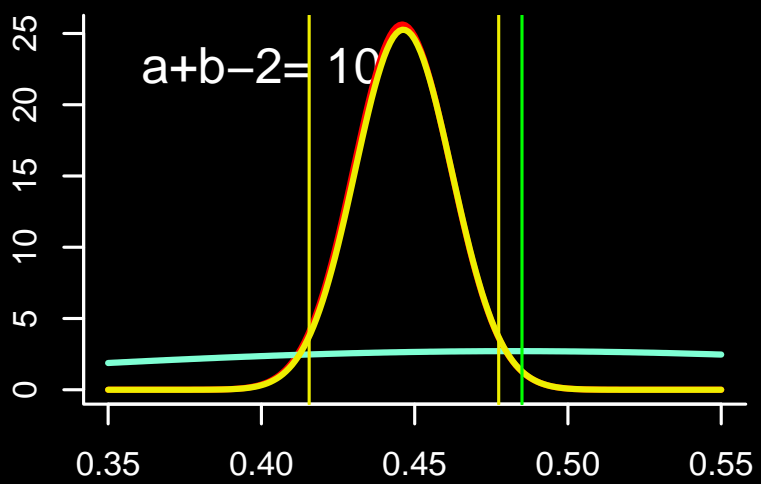
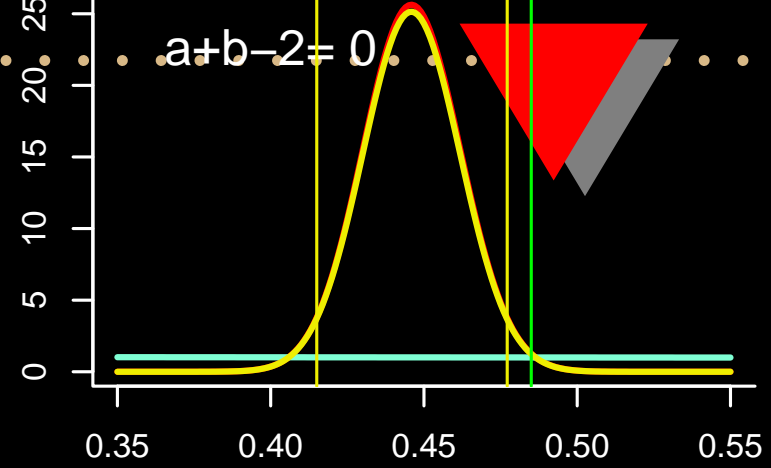
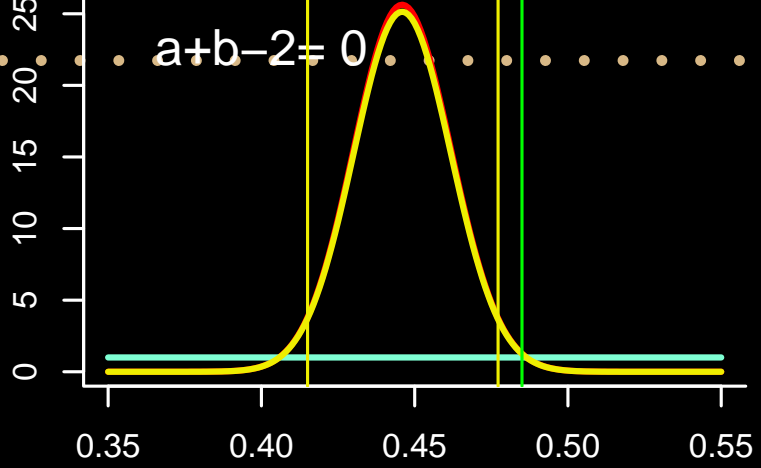
$$\propto \theta^{y+\alpha-1} (1 - \theta)^{n-y+\beta-1} = \text{Beta}(\alpha + y, \beta + n - y)$$

-> the BETA prior distribution is a **conjugate family** for
the BINOMIAL likelihood











How does posterior COMPROMISE between prior and the data?

- The compromise depends on how much weight prior has (or how much **informative** it is) w.r.t. the data at hand



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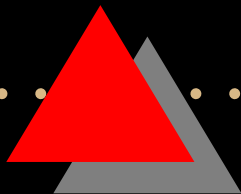
- The compromise depends on how much weight prior has (or how much **informative** it is) w.r.t. the data at hand
- i.e., in the binomial case, depends on the relative weight of

$$\alpha + \beta - 2$$

\approx number of *prior observations* (\sim **prior precision**)

Note: precision=1/variance, $\text{var} = \frac{\theta(1-\theta)}{\alpha+\beta+1}$

w.r.t. n , the sample size



A first SENSITIVITY ANALYSIS

concept of sensitivity: sensitivity or robustness of the inferences to the choice of the prior

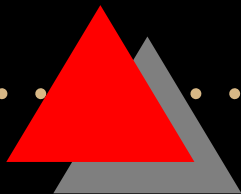
Prior information		Posterior information	
$\alpha + \beta - 2$	mean	mean	95% interval
0	0.500	0.446	[0.415 , 0.477]
0	0.485	0.446	[0.415 , 0.477]
10	0.485	0.446	[0.416 , 0.477]
100	0.485	0.450	[0.420 , 0.479]
1000	0.485	0.466	[0.444 , 0.488]
10000	0.485	0.482	[0.472 , 0.491]

NOTE: in placenta previa example $n \approx 1000$ and $\bar{y} = 0.446$

The SIMULATION-based estimation approach



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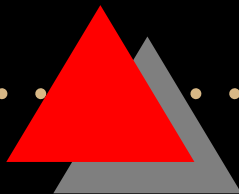
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- The modern approach to Bayesian estimation has become closely linked to **simulation-based estimation methods**.
- In fact, Bayesian estimation focuses on **estimating the entire density of a parameter**.
- This density estimation is based on generating samples from the posterior density of the parameters themselves or of functions of parameters.

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 - Then, direct calculations are feasible or **direct simulation** from it can be performed.
 - However, even if posterior density cannot be explicitly integrated, **iterative simulation methods** (or **MCMC**) are alternatively used. We will see them in future lab's.





a first (direct) simulation

Congdon book, pag. 31, sec. 2.11

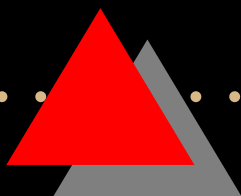
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- We are interested in assessing the probability that a randomly sampled adult would respond 'immoral'.
- In the inference we might use evidence from previous polls on the proportion of the population generally likely to consider a President's actions immoral.



a first (direct) simulation

The R code is in `betabin.r` at the course web page

- We present Bayesian inference about the probability of an adult responding ‘immoral’ assuming different Beta priors:

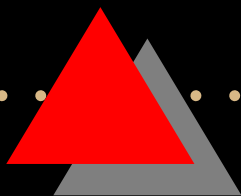
1. $\alpha = \beta = 1$ prior information ~ 0 $E = 1/2$

2. $\alpha = \beta = 0.001$ prior information < 0 $E = 1/2$

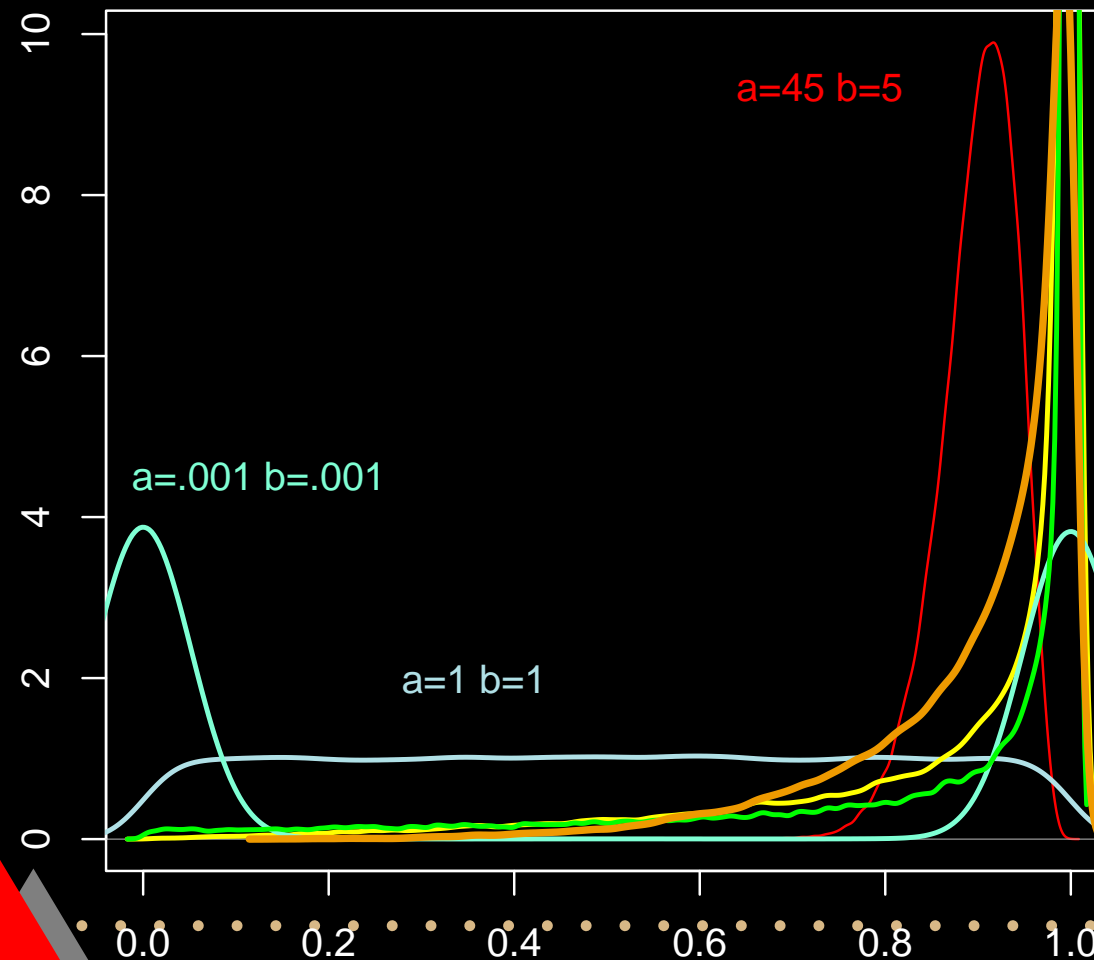
3. $\alpha = 1$ $\beta = 0.11$ prior information < 0 $E = 0.9$

4. $\alpha = 1.8$ $\beta = 0.2$ prior information ~ 0 $E = 0.9$

5. $\alpha = 4.5, 45$ $\beta = 0.5, 5$ prior information $\sim 5,50$
 $E = 0.9$



1., 2. are both **non informative**, but 2. is a reasonable choice for 'one-off' events (or for correlated data) 3., 4. may be assumed on the basis of previous polls. Although $E=0.9$ they still are **diffuse**. 5., 6 are **increasingly informative**.





Legend for the next figure →

in each figure:

- curves: **histogram** of 10,000 draws from the posterior $\text{Beta}(150+\alpha, 601+\beta)$; **likelihood** $\propto \text{Bin}(150, 751)$.
intervals: **Unif-Bin 95% posterior interval**; **95% (Beta(150+ α , 601+ β)) posterior interval**; **Normal approximation of the 95% posterior interval**; **Inverted 95% posterior interval on the logit scale**.



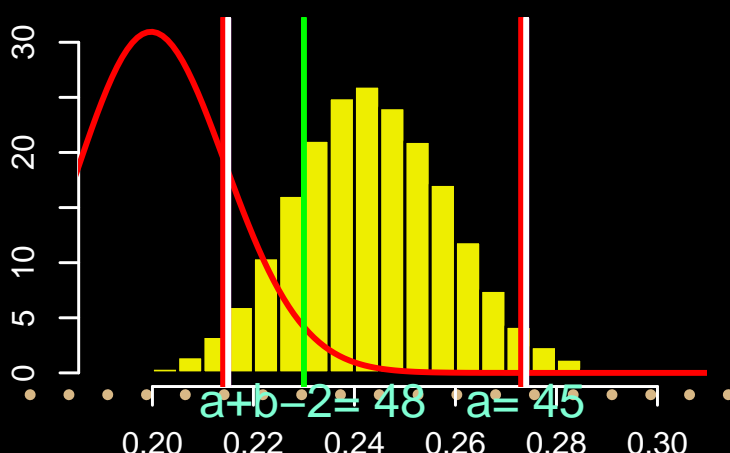
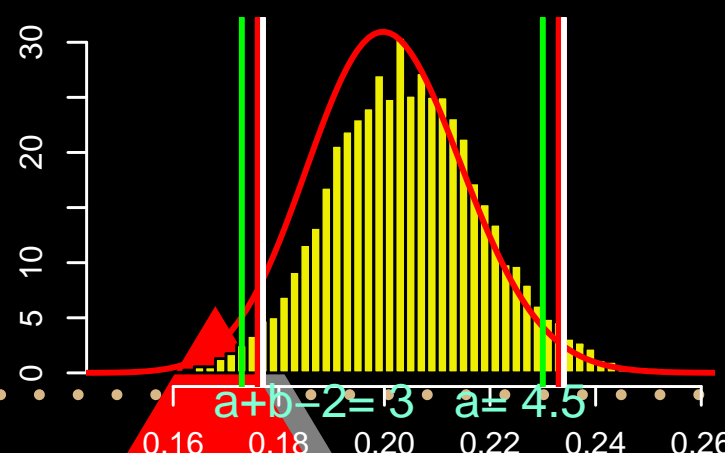
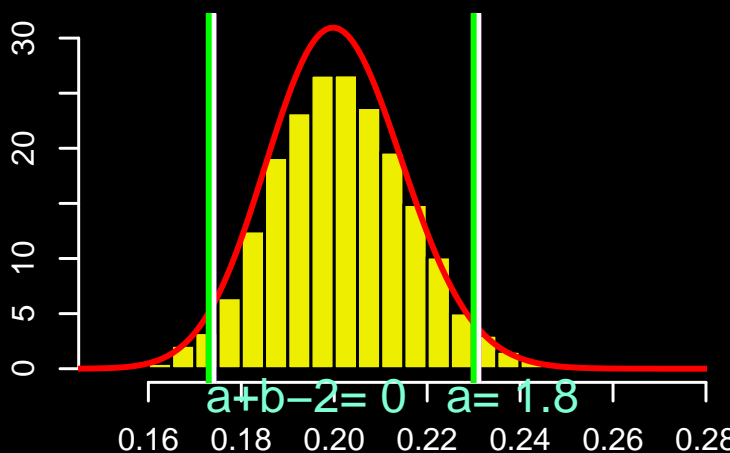
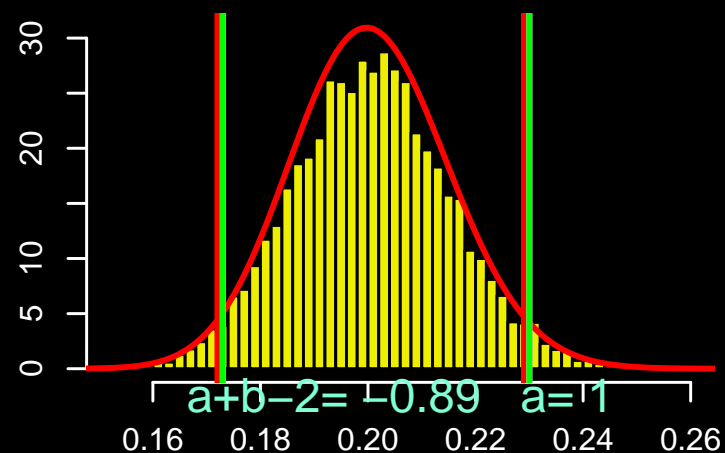
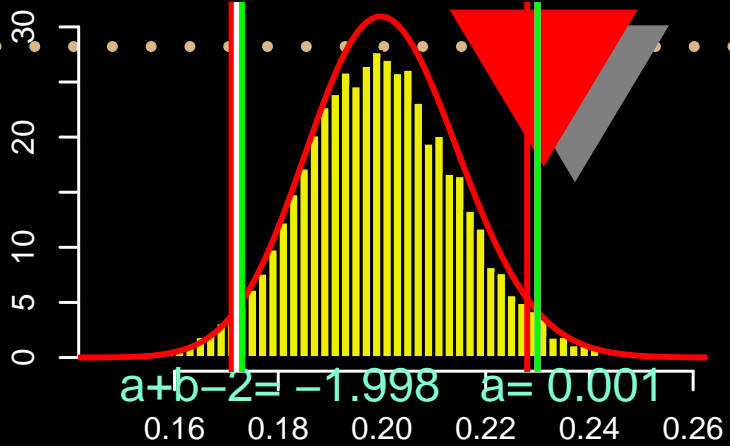
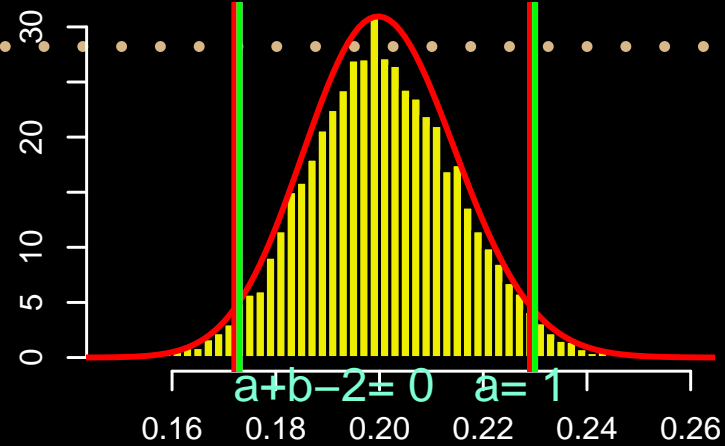
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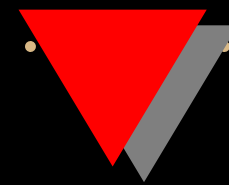
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- Though θ is close to 0, because of the **large sample size (751)**, the normal approximation is good as well as posterior inferences are *insensitive* to prior choice (even if discordant to data), at least for prior information

≤ 0 .







And what about if our sample size was only $n = 5$, with $y = 1$ adults considering immoral the President's actions? ->

NOTE: the empirical mean still is $y/n = 0.2$

