Wednesday 13 February 2002

Analysis of pigs data

Data: Body weights of 48 pigs at 9 successive follow-up visits.

This is an equally spaced data. It is always a good habit to reshape the data, so we can easily switch form wide to long or long to wide depending on the required analysis. The data is in the wide format; let's reshape it into long format.

. reshape groups time 1-9 ## Since we have observations at 9 time points

- . reshape vars week ## declare the variables
- . reshape cons Id ## declare constants
- . reshape long (wide)

While reshaping the data from any format, the first three steps are the same. The lat command, reshape long or reshape wide depends on the structure of the original data.

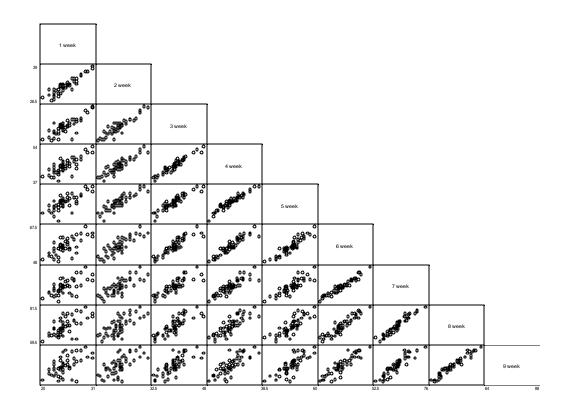
Now we have the data in long format, to go into the wide form, just type reshape wide. You don't have to repeat the first three commands.

Exploratory analysis

Lets just make some scatter plots of the data. First we plot the scatter plot matrix. For this we require the data in wide format.

. reshape wide

. graph week1 week2 week3 week4 week5 week6 week7 week8 week9, matrix half

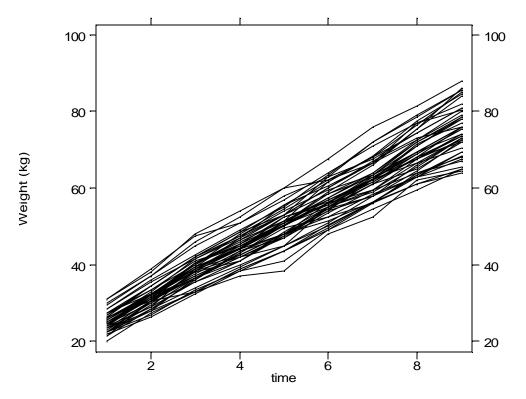


What do you conclude from the above scatter plot matrix?

To see if the pigs gained weight over time lets plot the line (spaghetti) plot. For this we need the data in the long form.

- . reshape long
- . sort Id time

. graph week time, c(L) s(i) xlab(2 4 6 8) ylab rlab



What do you conclude from the graph?

The above figure is enough to explore the growth data. It is hard to pick out individual response profiles. We can add a second display, obtained form first standardizing each observation. This is achieved by, subtracting the mean, and dividing by the standard deviation of the 48 observations at each time (week). For this we would need the data in wide format.

[.] reshape wide

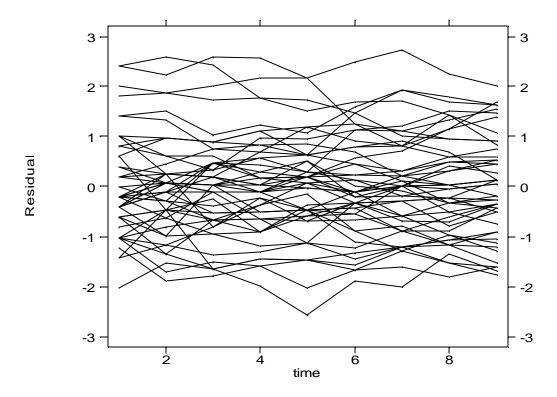
Now do the following for each of week1, week2, ..., week9 . sum week1

Variable	Obs	Mean	Std. Dev.	Min	Max
+					
week1	48	25.02083	2.468866	20	31
. gen Sweekl =				20	51

After you do this, we will have 9 new variables. To make the plot, again reshape long.

. sort Id time

. graph Sweek time, c(L) s(.) xlab(2 4 6 8) ylab(-3 -2 -1 0 1 2 3) rlab(-3 -2 -1 0 1 2 3)



The plot is able to highlight the degree of *tracking*, animals tend to maintain their relative size over time.

Exploring the correlation structure

Auto-correlation function. For this we require the data in long format.

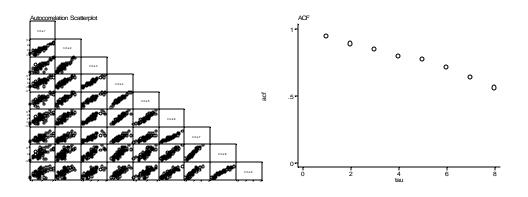
. autocor week time Id

Table 1

	timel	time2	time3	time4	time5	time6	time7	time8	time9
time1 time2 time3 time4 time5 time6 time7 time8	1.0000 0.9156 0.8015 0.7958 0.7494 0.7051 0.6551 0.6255	1.0000 0.9118 0.9084 0.8809 0.8353 0.7759 0.7133	1.0000 0.9582 0.9280 0.9058 0.8435 0.8167	1.0000 0.9621 0.9327 0.8681 0.8293	1.0000 0.9219 0.8546 0.8104	1.0000	1.00	00	.0000
time9	0.5581	0.6638	0.7689	0.7856	0.7856	0.8893	0.9170 	0.969	5 1.0000

Table 2

	acf
1.	.9425781
2.	.8870165
3.	.8462396
4.	.7962576
5.	.7724156
6.	.7121489
7.	.6407955
8.	.5581002



Notice from the main diagonal of the scatter plot matrix there is positive correlation between repeated observations on the same animal that are 1 week apart. The degree of correlation decreases as the observations are moved farther from the diagonals. Also the correlation is reasonably consistent along the diagonal in the matrix. This indicates that the correlation depends on the time between observations than their absolute times. The estimated correlation matrix for this data is given in table 1. The correlations show some tendency of decrease with increasing time lag. Assuming stationarity, a single correlation estimate can be obtained for each distinct value of the time separation or lag, $|t_{ij} - t_{ik}|$. This corresponds to pulling observation pairs along the diagonals of the scatter plot matrix. The autocorrelation function takes the value as in table 2.

Before we proceed with the analysis, lets look at some theory.

 $y_{ij} j = 1, 2, ..., n$ be the sequence of observed measurements on the *ith* of the *m* subjects and $t_j, j = 1, 2, ..., n$ be the corresponding times at which the measurements are taken on each unit. Associated with each y_{ij} are the values, $x_{ijk}, k = 1, 2, ..., p$ of *p* explanatory variables. We assume that y_{ij} are realizations of random variables Y_{ij} which follow the regression model

$$Y_{ij} = \boldsymbol{b}_1 x_{ij1} + \dots + \boldsymbol{b}_p x_{ijp} + \boldsymbol{e}_{ij},$$

In the classical linear model we assume the errors to be mutually independent normal random variables. In our context, the longitudinal structure of the data means that we expect the errors to be correlated within subjects.

Let $y_i = (y_{i1}, y_{i2}, ..., y_{in})$ be the observed sequence of measurements on the *ith* subject and $y = (y_1, y_2, ..., y_m)$ be the complete set of N = nm observations. Let **X** be the matrix of explanatory variables.

$$Y \sim MVN(X\boldsymbol{b},\boldsymbol{s}^2 V)$$

The Uniform Correlation Model

In this model we assume that there is positive correlation between any two measurements.

The Exponential Correlation Model

The correlation between a pair of measurements on the same unit decays towards zero as the time separation between measurements increases.

The exponential correlation model is sometimes called the *first order autoregressive model*.

For the data on pigs we fit a couple of models. We require the data in long format

1. Ordinary least squares ignoring correlation

.regress week time

Source	SS	df MS	5		Number of $obs = 432$ F(1, 430) = 5757.41
Model Residual	111060.882 8294.72677		50.882 900622		Prob > F = 0.0000 R-squared = 0.9305 Adj R-squared = 0.9303
Total	119355.609	431 276.9	927167		Root MSE = 4.392
week	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
time _cons	6.209896 19.35561	.0818409 .4605447	75.878 42.028	0.000	6.049038 6.370754 18.45041 20.26081

2. Independent correlation model

. xtgee week time, i(Id) corr(ind)

Iteration 1: tolerance = 1.848e-15

GEE popula	tion-averaged	model		Number of	obs =	432
Group vari	.able:		Id	Number of	groups =	48
Link:			identity	Obs per g	roup: min =	9
Family:			Gaussian		avg =	9.0
Correlatic	on:	in	dependent		max =	9
			-	Wald chi2	(1) =	5757.41
Scale para	meter:		19.29006	Prob > ch	i2 =	0.0000
Pearson chi2(430):		8294.73	Deviance	=	8294.73	
Dispersion	(Pearson):		19.29006	Dispersio	n =	19.29006
week	Coef.	Std. Err.	Z	P> z	[95% Conf.	Interval]
+ time	6.209896	.0818409	 75.878	0.000	6.049491	6.370301
_cons	19.35561	.4605447	42.028	0.000	18.45296	20.25826

xtcorr ## estimates the correlation matrix

Estimated within-Id correlation matrix R:

	c1	c2	c3	c4	c5	сб	с7	c8	с9
r1	1.0000								
r2	0.0000	1.0000							
r3	0.0000	0.0000	1.0000						
r4	0.0000	0.0000	0.0000	1.0000					
r5	0.0000	0.0000	0.0000	0.0000	1.0000				
rб	0.0000	0.0000	0.0000	0.0000	0.0000	1.0000			
r7	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000	1.0000		
r8	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000	1.0000	
r9	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000	1.0000

3. Exponential correlation model

. xtgee week time, i(Id) corr(arl) t(time) Iteration 1: tolerance = .02516015 Iteration 2: tolerance = .00009265 Iteration 3: tolerance = 4.393e-07Number of obs=Id timeNumber of groups=identityObs per group: min =Gaussian 432 GEE population-averaged model Group and time vars: 48 Link: 9 avg = 9.0 Family: Gaussian Correlation: 9 AR(1) max = Wald chi2(1) = 6255.06 Prob > chi2 = 0.0000 19.35735 Scale parameter: _____ week | Coef. Std. Err. z P>|z| [95% Conf. Interval] _____+ time6.272119.079304779.0890.0006.1166856.427553_cons18.84162.675981927.8730.00017.5167220.16652 _____

. xtcorr

Estimated within-Id correlation matrix R:

	c1	c2	c3	с4	c5	сб	с7	с8	с9
r1	1.0000								
r2	0.9172	1.0000							
r3	0.8413	0.9172	1.0000						
r4	0.7716	0.8413	0.9172	1.0000					
r5	0.7077	0.7716	0.8413	0.9172	1.0000				
rб	0.6491	0.7077	0.7716	0.8413	0.9172	1.0000			
r7	0.5954	0.6491	0.7077	0.7716	0.8413	0.9172	1.0000		
r8	0.5461	0.5954	0.6491	0.7077	0.7716	0.8413	0.9172	1.0000	
r9	0.5008	0.5461	0.5954	0.6491	0.7077	0.7716	0.8413	0.9172	1.0000

4. Between effects WLS (usually done in case of unbalanced data instead of OLS)

. xtreg week time, be wls i(Id)

5. Random-effects model

. xtreg week time, re i(Id)

	fects GLS regr iable (i) : Id			obs = groups =		
bet	chin = 0.9851 cween = 0.0000 erall = 0.9305)	Obs per gi	roup: min = avg = max =	9 9.0 9	
	fects u_i ~ Ga X) = O			Wald chi2 Prob > ch:	(1) = i2 =	
	Coef.		Z	P> z	[95% Conf.	Interval]
time	6.209896 19.35561	.0390633				
sigma_u sigma_e rho	2.0963559	(fraction	of variance	due to u_i)		

6. xtreg, mle

. xtreg week time, i(id) mle

	ects ML regre able (i) : Io			obs = groups =		
Random eff	ects u_i ~ Ga	aussian		Obs per g	roup: min = avg = max =	9.0
Log likeli	.hood = -1014	1.9268) = i2 =	
	Coef.				[95% Conf.	Interval]
time	6.209896 19.35561	.0390124 .5974055	159.178	0.000 0.000		20.52651
/sigma_e	3.84935 2.093625	.4058114 .0755471	9.486	0.000	3.053974 1.945555	4.644725
	.771714				.6876303	
Likelihood	l ratio test d	of sigma_u=0:	chi2(1)	= 472.65	Prob > chi	2 = 0.0000

Note that, xtreg, re and xtreg, mle give the same estimates.

The following table might be useful for future reference. To learn more explore the xt set of commands in STATA

If I want to do	STATA command
OLS	regress Y X
Between-effects model (WLS)	xtreg Y X, be i(id) wls if data is unbalanced specify wls instead of OLS, default is OLS
Random-effects model (mixed)	xtreg Y X, re i(id)
Fixed effect model (within)	xtreg Y X, fe i(id)
MLE random effect model	xtreg Y X, i(id) mle
Population average model	<pre>xtreg Y X, pa i(id) This command is equivalent to xtgee Y X, f(gaussian) link(id) corr(exc)</pre>
Generalized least squares	<pre>xtgls Y X, i(id) corr(arl or ind) See STATA help before you use this command</pre>
Population averaged panel data using GEE	<pre>xtgee Y X, f(family) l(link) corr(correlation) i(id) for example: xtgee Y X1 X2, f(gauss) l(id) corr(exc) i(id) would estimate population-averaged linear regression with equal correlation comparable to random-effects regression fro xtreg</pre>