Assessing variants in the human genome

Ingo Ruczinski

Department of Biostatistics

Johns Hopkins Bloomberg School of Public Health

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Very large data sets



Very large data sets



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http://biostat.jhsph.edu/~iruczins/ingo@jhu.edu

Acknowledgments

Collaborators: Kathleen Barnes, Terri Beaty, Benilton

Carvalho, Bob Cole, Rafael Irizarry, Tom Louis, Rasika Mathias, Matt Ritchie, Rob Scharpf,

Holger Schwender, Keith West.

Computing support: Marvin Newhouse, Jiong Yang.

Funding: NIH R01 DK061662, GM083084, HL090577,

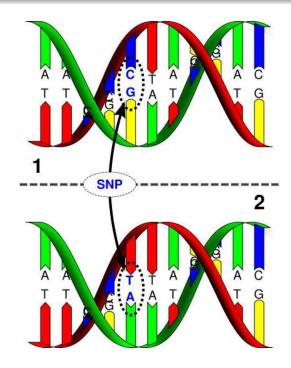
and a CTSA grant to the Johns Hopkins

Medical Institutions.

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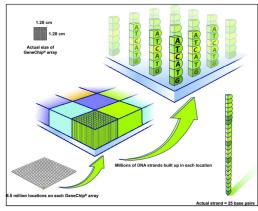
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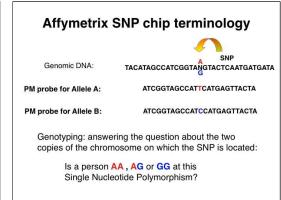
Single nucleotide polymorphisms



urgi.versailles.inra.fr

Genomic arrays



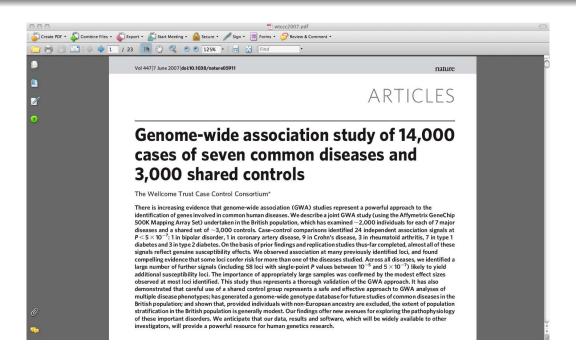


http://www.affymetrix.com

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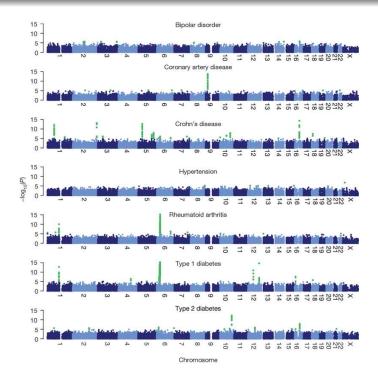
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WTCCC



Wellcome Trust Case Control Consortium (2007). Nature 447(7145): 661-78.

WTCCC

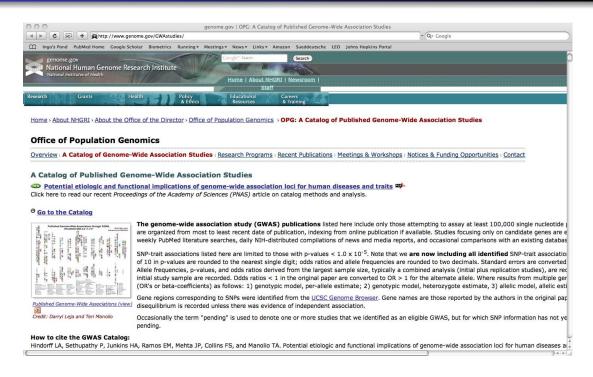


Wellcome Trust Case Control Consortium (2007). Nature 447(7145): 661-78.

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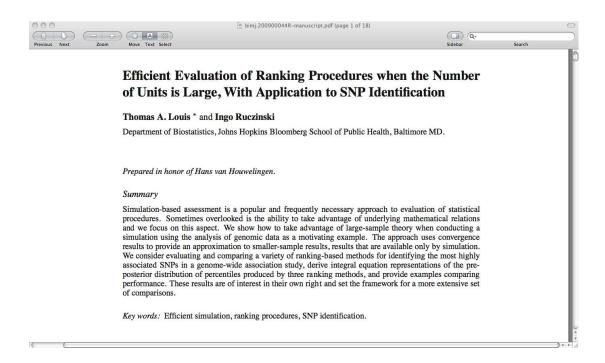
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Results



http://www.genome.gov/GWAstudies/

Ranking



[Lou · Ruc | Biom · J 2010] ● [Ruc · · · Lou | Bayes · Stat 2010]

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Case-parent trios

Recruitment Site	CL	CLP	CP	Total
Utah	68	96	52	216
Norway	106	174	107	387
Korea	19	40	5	64
Maryland	19	71	25	115
Pittsburgh	26	70	11	107
Singapore	15	45	53	113
Taiwan	42	176	74	292
Iowa	16	29	24	69
Denmark	6	15	5	26
Philippines	0	94	0	94
WuHan	39	136	42	217
Shandong Province	54	129	30	213
Western China	43	63	38	144
Total	453	1138	466	2057

Case-parent trios

$$F:12 \longrightarrow M:12$$
 $F:12 \longrightarrow M:12$ $F:12 \longrightarrow M:12$

$$F:11 \longrightarrow M:12$$
 $F:11 \longrightarrow M:12$

C:11

$$F: 12 \longrightarrow M: 22$$
 $F: 12 \longrightarrow M: 22$

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Genotypic TDT

Assume that at a certain locus the father has alleles 11 and the mother has alleles 12. The four *Mendelian children* thus have alleles 11, 12, 11, and 12.

Assume the affected proband has genotype 11.

The three *Pseudo controls* then have the genotypes 11, 12, and 12.

	Υ	Χ
Affected proband	1	11
Pseudo control #1	0	11
Pseudo control #2	0	12
Pseudo control #3	0	12

We can use coonditional logistic regression to analyze the data.

Allelic TDT

The transmission disequilibrium test measures the over-transmission of an allele from parents to affected offsprings. For a set of n parents with alleles 1 and 2 at a genetic locus, each parent can be summarized by the transmitted and the non-transmitted allele:

		Non-TA		
		1	2	Σ
≰	1	a	b	a + b
	2	С	d	c + d
Σ		a + c	b + d	2n

Only the heterozygous parents contribute information!

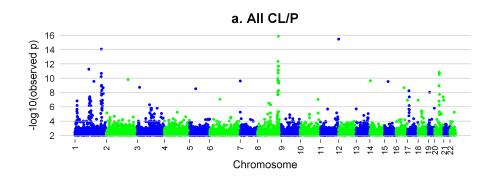
Under the null of no association, $\frac{(b-c)^2}{b+c} \sim \chi_1^2$

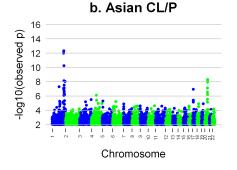
 \rightarrow Even better, use binom.test() in R.

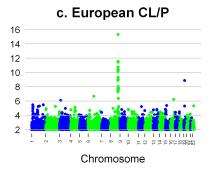
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GWAs results

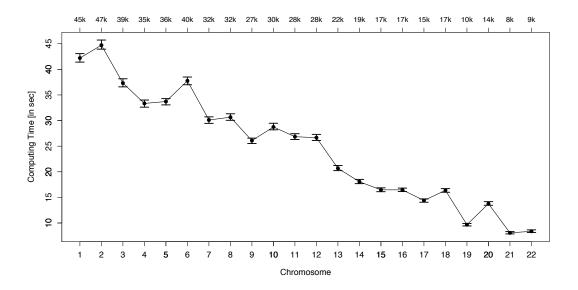






[BEA · · · RUC · · · SCO | NAT-GEN 2010]

Fast genotypic TDT



[SCH · · · Ruc | TECH · REP 2010]

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Candidate genes

OPEN @ ACCESS Freely available online



Genetic Determinants of Facial Clefting: Analysis of 357 Candidate Genes Using Two National Cleft Studies from Scandinavia

Astanand Jugessur^{1,9}, Min Shi^{2,9}, Håkon Kristian Gjessing^{3,4}, Rolv Terje Lie^{4,5}, Allen James Wilcox⁶, Clarice Ring Weinberg², Kaare Christensen⁷, Abee Lowman Boyles⁶, Sandra Daack-Hirsch⁸, Truc Nguyen Trung⁵, Camilla Bille⁷, Andrew Carl Lidral⁹, Jeffrey Clark Murray^{7,9}*

1 Craniofacial Development, Musculoskeletal Disorders, Murdoch Children's Research Institute, Royal Children's Hospital, Parkville, Australia, 2 Biostatistics Branch, National Institute of Environmental Health Sciences (NIEHS), Research Triangle Park, Durham, North Carolina, United States of America, 3 Department of Epidemiology (EPAM), Norwegian Institute of Public Health, Oslo, Norway, 4 Section for Epidemiology and Medical Statistics, Department of Public Health and Primary Health Care, University of Bergen, Bergen, Norway, 5 Medical Birth Registry of Norway, Norwegian Institute of Public Health, Bergen, Norway, 6 Epidemiology Branch, National Institute of Environmental Health Sciences (NIEHS), Research Triangle Park, Durham, North Carolina, United States of America, 7 Department of Epidemiology, University of Southern Demark, Oderse, Demark, Bollege of Nursing, University of lowa City, Iowa City, Iowa, United States of America, 9 Departments of Pediatrics, Epidemiology and Biological Sciences, University of Iowa, United States of America

Abstract

Background: Facial clefts are common birth defects with a strong genetic component. To identify fetal genetic risk factors for clefting, 1536 SNPs in 357 candidate genes were genotyped in two population-based samples from Scandinavia (Norway: 562 case-parent and 592 control-parent triads; Denmark: 235 case-parent triads).

Methodology/Principal Findings: We used two complementary statistical methods, TRIMM and HAPLIN, to look for associations across these two national samples. TRIMM tests for association in each gene by using multi-SNP genotypes from case-parent triads directly without the need to infer haplotypes. HAPLIN on the other hand estimates the full haplotype distribution over a set of SNPs and estimates relative risks associated with each haplotype. For isolated cleft lip with or without cleft palate (I-CLP), TRIMM and HAPLIN both identified significant associations with IRF6 and ADH1C in both populations, but only HAPLIN found an association with FGF12. For isolated cleft palate (I-CP), TRIMM found associations with AUX3, MCX, and PDGFC in both populations, but only the association with PDGFC was identified by HAPLIN. In addition, HAPLIN identified an association with ETV5 that was not detected by TRIMM.

Conclusion/Significance: Strong associations with seven genes were replicated in the Scandinavian samples and our approach effectively replicated the strongest previously known association in clefting—with IRF6. Based on two national cleft cohorts of similar ancestry, two robust statistical methods and a large panel of SNPs in the most promising cleft candidate genes to date, this study identified a previously unknown association with clefting for ADHTC and provides additional candidates and analytic approaches to advance the field.

Citation: Jugessur A, Shi M, Gjessing HK, Lie RT, Wilcox AJ, et al. (2009) Genetic Determinants of Facial Clefting: Analysis of 357 Candidate Genes Using Two National Cleft Studies from Scandinavia. PLoS ONE 4(4): e5385. doi:10.1371/journal.pone.0005385

Parent-of-origin effects

			I	Paternal	20	40	N	Maternal	201		
	<u>.</u>	TAT				TAT			PO-LRT ^b		
No.	SNP name	T	NT	P-value	ORc	T	NT	P-value	OR ^c	OR ^d	P-value
1	rs7771980	9	8	0.808	1.13	16	18	0.732	0.89	0.79	0.692
2	rs2677104	25	30	0.500	0.83	22	24	0.768	0.92	1.10	0.811
3	rs2819855	36	34	0.811	1.06	37	25	0.128	1.48	1.40	0.342
4	rs2819854	35	36	0.906	0.97	37	29	0.325	1.28	1.32	0.417
5	rs910586	15	13	0.705	1.15	20	5	0.003	4.00	3.59	0.036
6	rs2819853	14	12	0.695	1.17	18	5	0.007	3.60	3.19	0.063
7	rs765724	15	13	0.705	1.15	20	6	0.006	3.33	2.97	0.065
8	rs1343799	14	12	0.695	1.17	18	5	0.007	3.60	3.19	0.063
9	rs2819861	13	12	0.841	1.08	19	5	0.004	3.80	3.73	0.036
10	rs2790103	16	11	0.336	1.45	20	5	0.003	4.00	2.86	0.092
11	rs2790093	15	12	0.564	1.25	18	5	0.007	3.60	2.99	0.079
12	rs2790098	15	12	0.564	1.25	19	6	0.009	3.17	2.60	0.110
13	rs4714854	15	12	0.564	1.25	19	6	0.009	3.17	2.60	0.110
14	rs9472494	15	14	0.853	1.07	22	7	0.005	3.14	2.99	0.051
15	rs2396442	17	14	0.590	1.21	24	8	0.005	3.00	2.51	0.086
16	rs1934328	41	17	0.002	2.41	35	33	0.808	1.06	0.44	0.029
17	rs7773875	33	21	0.102	1.57	32	32	1.000	1.00	0.65	0.245
18	rs7771889	36	18	0.014	2.00	40	31	0.285	1.29	0.64	0.238
19	rs10485422	15	13	0.705	1.15	17	6	0.022	2.83	2.42	0.135
20	rs6904353	13	14	0.847	0.93	18	11	0.194	1.64	1.78	0.294
21	rs13207392	16	15	0.857	1.07	19	7	0.019	2.71	2.50	0.102
22	rs7748231	13	13	1.000	1.00	18	11	0.194	1.64	1.64	0.373
23	rs10948237	13	14	0.847	0.93	18	11	0.194	1.64	1.78	0.294
24	rs1928533	12	13	0.841	0.92	15	13	0.705	1.15	1.27	0.671

[SUL · · · RUC · · · BEA | GEN-EPI 2008]

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Copy number estimates are noisy

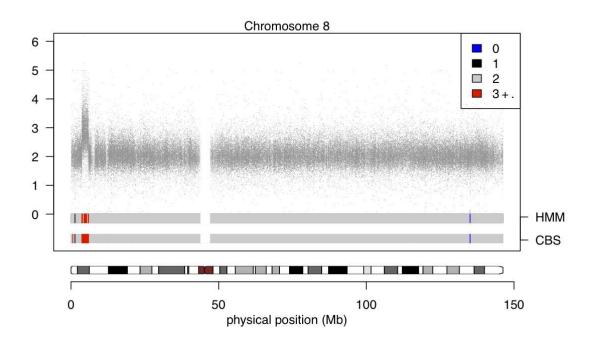
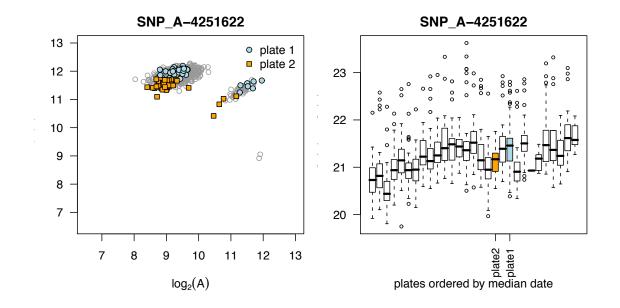


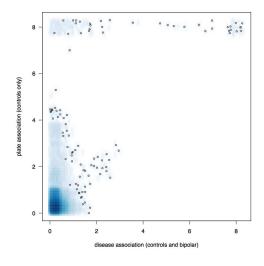
Plate effects

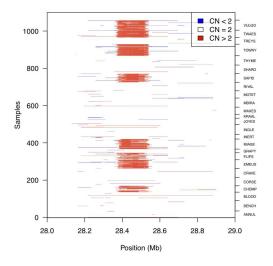


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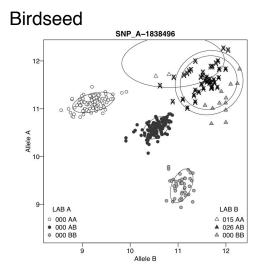
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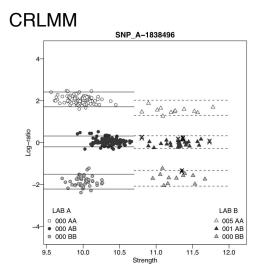
Confounding of plate and disease





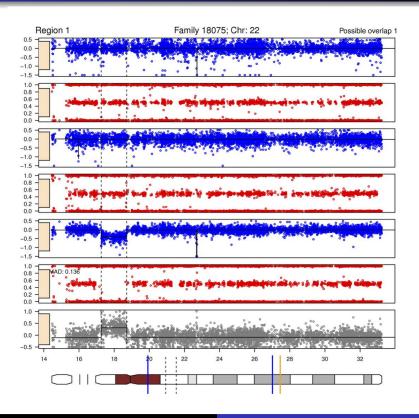
Genotype estimates are more robust



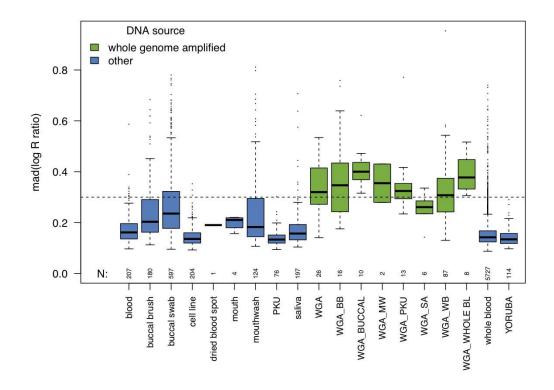


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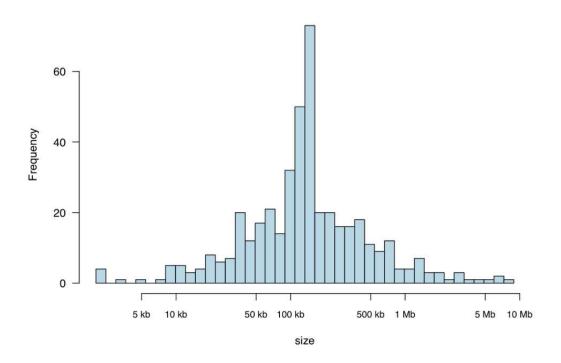


DNA source

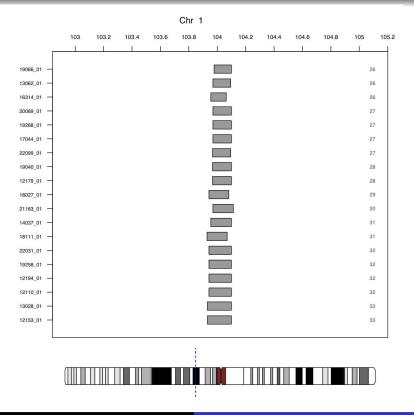


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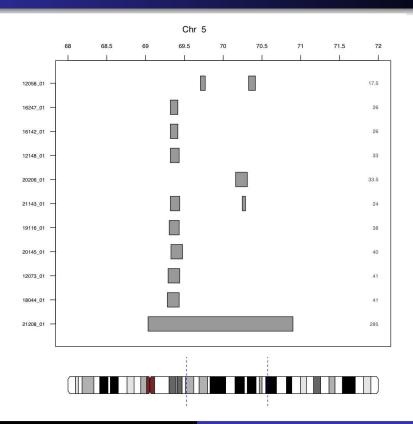


De-novo deletions

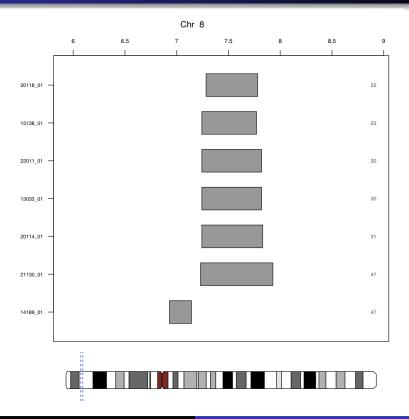


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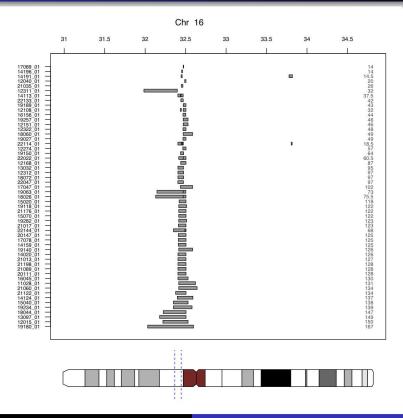


De-novo deletions

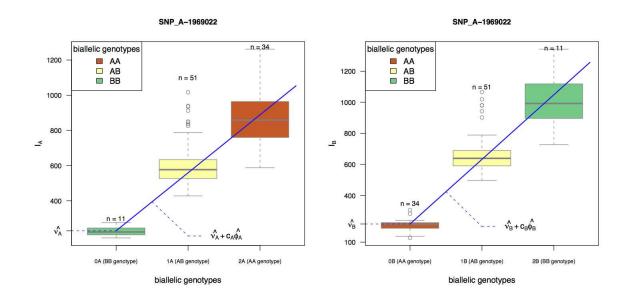


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Allele specific copy numbers



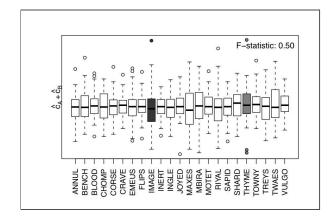
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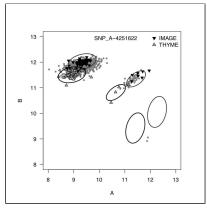
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Allele specific copy numbers

At locus i, for subject j in plate p, we have for allele $k \in \{A, B\}$

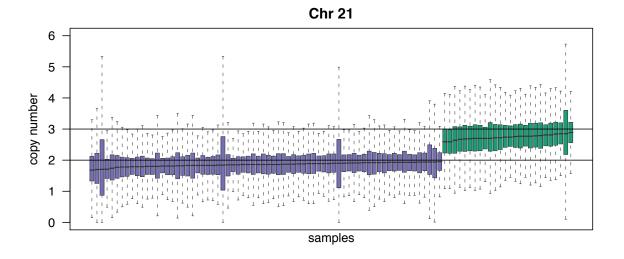
$$I_{\mathit{kijp}} =
u_{\mathit{kip}} \delta_{\mathit{kijp}} + \phi_{\mathit{kip}} c_{\mathit{kijp}} \epsilon_{\mathit{kijp}} \implies \hat{c}_{\mathit{kijp}} = \max \left\{ \frac{1}{\hat{\phi}_{\mathit{kip}}} \left(I_{\mathit{kijp}} - \hat{\nu}_{\mathit{kip}} \right), \ 0 \right\}$$





[SCH · IRI · RIT · CAR · RUC | TECH·REP 2010] • [SCH · RUC · CAR · DOA · CHA · IRI | BIOSTAT 2010]

Trisomy 21

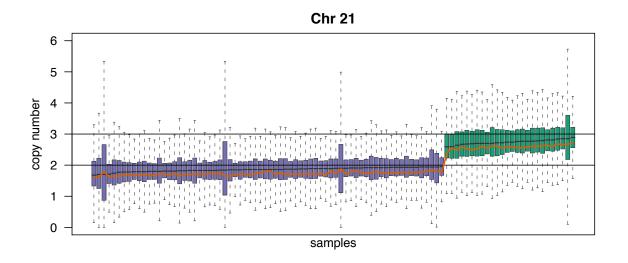


Samples from Aravinda Chakravarti and Betty Doan

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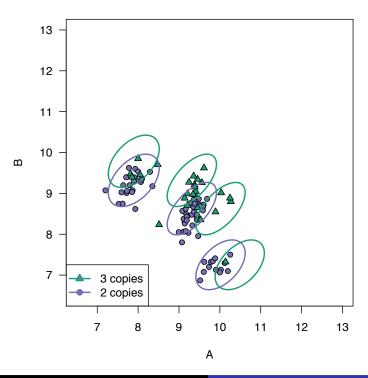
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Trisomy 21



Samples from Aravinda Chakravarti and Betty Doan

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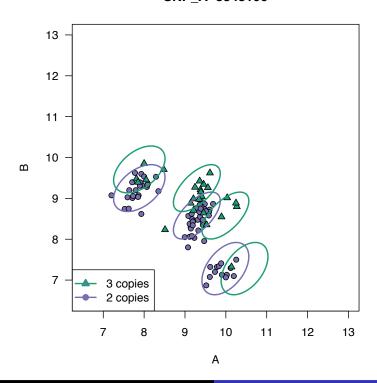


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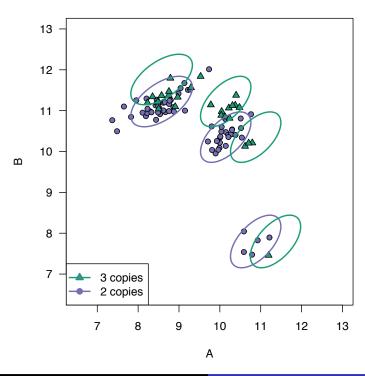
A versus B plots

SNP_A-8348190



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SNP_A-8341330

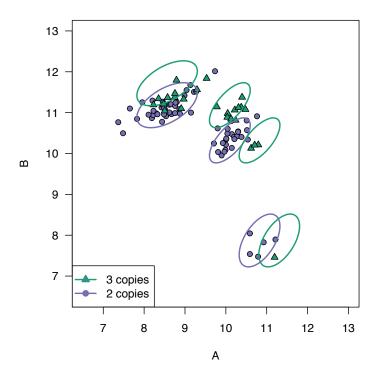


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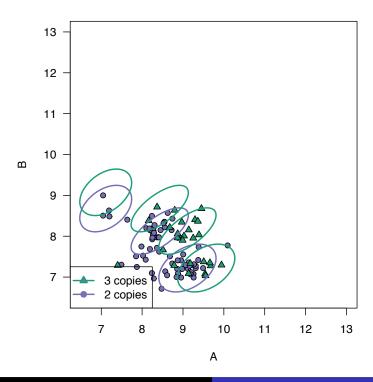
A versus B plots

SNP_A-8341330



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SNP_A-8339372

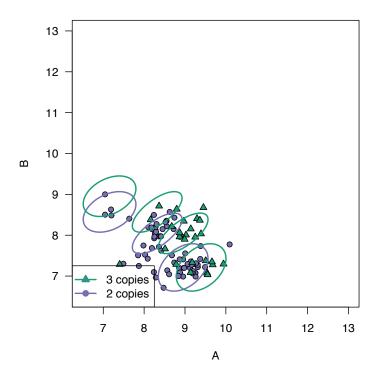


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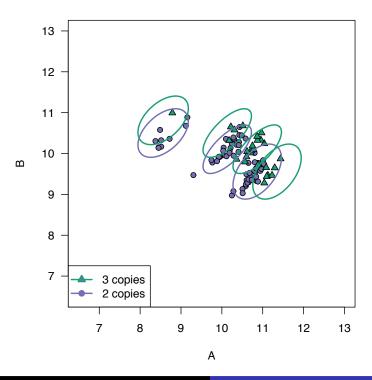
A versus B plots

SNP_A-8339372



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SNP_A-8340560

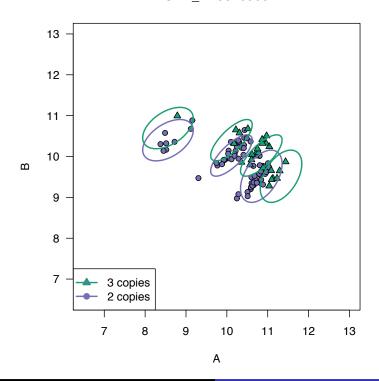


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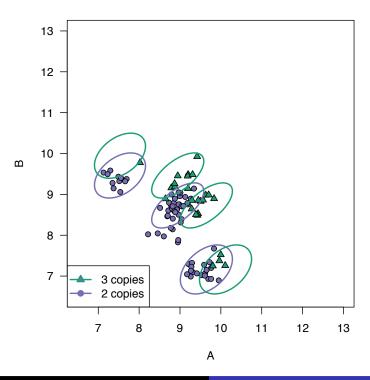
A versus B plots

SNP_A-8340560



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SNP_A-1969323

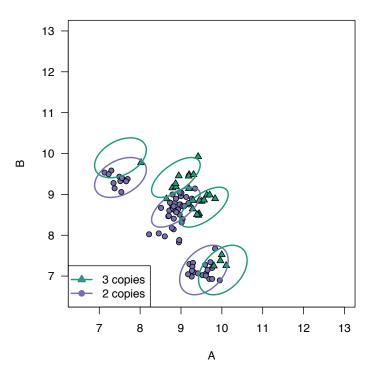


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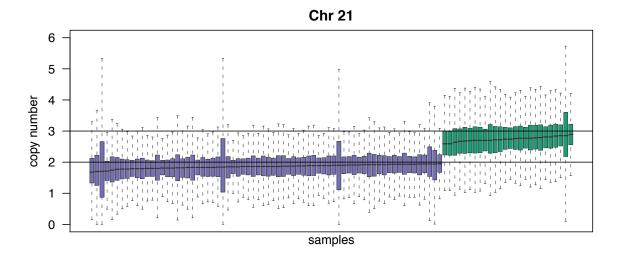
A versus B plots

SNP_A-1969323



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Trisomy 21

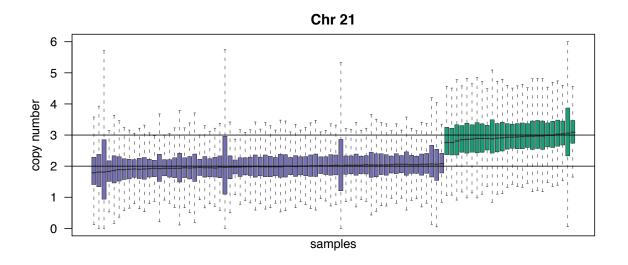


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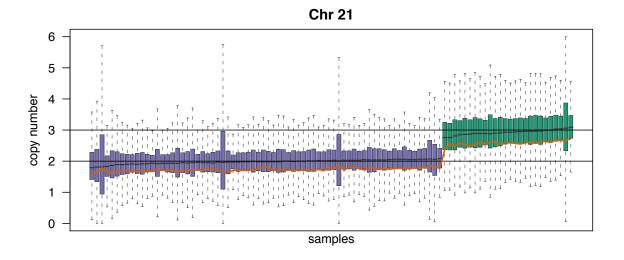
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Trisomy 21



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Trisomy 21

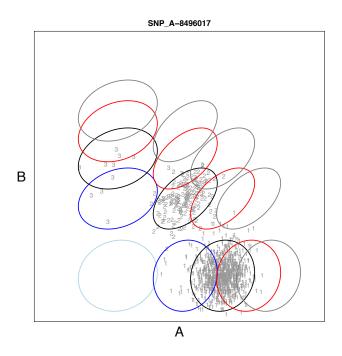


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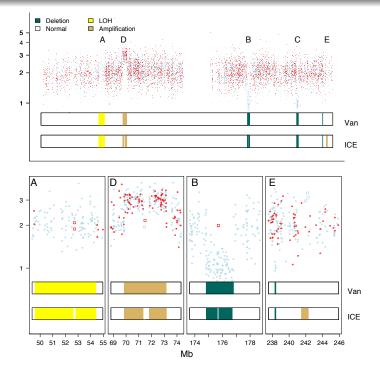
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Prediction regions for copy number



[Sch \cdot Ruc \cdot Car \cdot Doa \cdot Cha \cdot Iri | Biostat 2010]

Vanilla and ICE HMMs for genotype and copy number

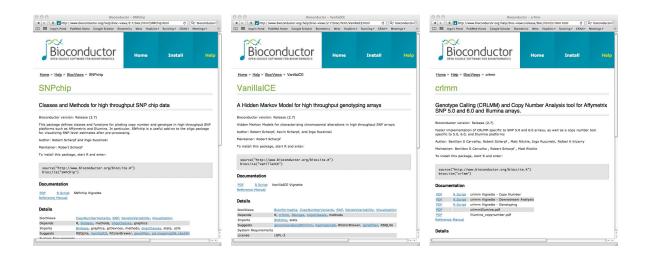


[Sch · Par · Pev · Ruc | Aoas 2008]

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Open source software



[Sch · · · · Ruc | Bioinf 2007] ● [Sch · Ruc | M·Mol·Bio 2010] ● [Sch · Ruc · · · Iri | Biostat 2010]

A software vignette

Using the R Package crlmm for Genotyping and Copy Number Estimation

Robert B Scharpf Johns Hopkins University Rafael A Irizarry Johns Hopkins University

Matthew E Ritchie

Walter+Eliza Hall Institute of Medical Research

Benilton Carvalho University of Cambridge Ingo Ruczinski Johns Hopkins University

A botnoot

Genotyping platforms such as Affymetrix can be used to assess genotype-phenotype as well as copy number-phenotype associations at millions of markers. While genotyping algorithms are largely concordant when assessed on HapMap samples, tools to assesses copy number changes are more variable and often discordant. One explanation for the discordance is that copy number estimates are susceptible to systematic differences between groups of samples that were processed at different times or by different labs. Analysis algorithms that do not adjust for batch effects are prone to spurious measures of association. The R package crimm implements a multilevel model that adjusts for batch effects and provides allele-specific copy number. This paper illustrates a workflow for the estimation of allele-specific copy number, develops marker-and study-level summaries of batch effects, and demonstrates how the marker-level estimates can be integrated with complimentary Bioconductor software for inferring regions of copy number gain or loss. All analyses are performed in the statistical environment R. A compendium for reproducing the analysis is available from the author's website (http://www.biostat.jhsph.edu/~rscharpf/crlmmCompendium/index.html).

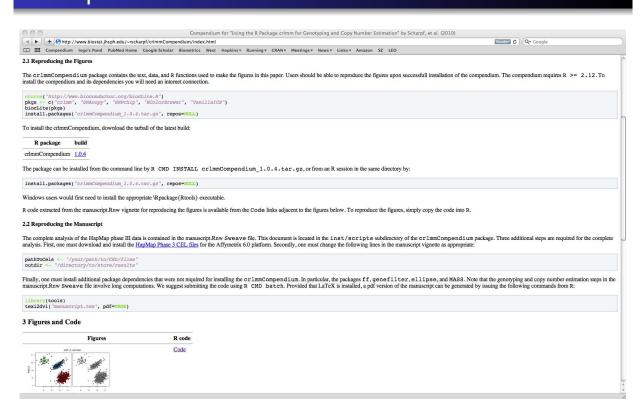
Keywords: copy number, batch effects, robust, multilevel model, high-throughput, oligonucleotide array.

[SCH · IRI · RIT · CAR · RUC | TECH·REP 2010]

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Compendium



Missing heritability



[MANOLIO ET AL | NATURE 461: 2009]

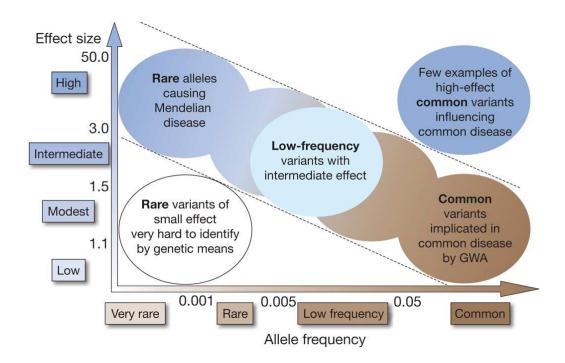
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Missing heritability

Disease	Number of loci	Proportion of heritability explained	Heritability measure
Age-related macular degeneration ⁷²	5	50%	Sibling recurrence risk
Crohn's disease ²¹	32	20%	Genetic risk (liability)
Systemic lupus erythematosus ⁷³	6	15%	Sibling recurrence risk
Type 2 diabetes ⁷⁴	18	6%	Sibling recurrence risk
HDL cholesterol ⁷⁵	7	5.2%	Residual phenotypic variance
Height ¹⁵	40	5%	Phenotypic variance
Early onset myocardial infarction ⁷⁶	9	2.8%	Phenotypic variance
Fasting glucose ⁷⁷	4	1.5%	Phenotypic variance

Missing heritability



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Genetic heterogeneity





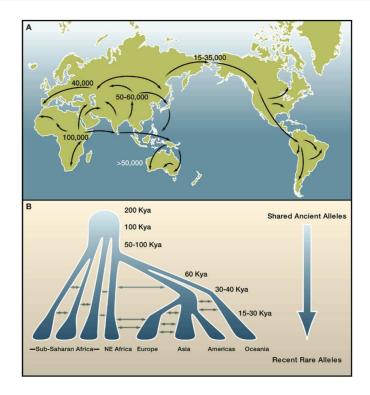
Genetic Heterogeneity in Human Disease

Jon McClellan^{1,*} and Mary-Claire King^{2,*}
¹Department of Psychiatry
²Departments of Medicine and Genome Sciences
University of Washington, Seattle, WA 98195-7720, USA
**Correspondence: drjack@uw.edu (J.M.), mcking@uw.edu (M.-C.K.)
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Strong evidence suggests that rare mutations of severe effect are responsible for a substantial portion of complex human disease. Evolutionary forces generate vast genetic heterogeneity in human illness by introducing many new variants in each generation. Current sequencing technologies offer the possibility of finding rare disease-causing mutations and the genes that harbor them.

[McClellan and King | Cell 141: 2010]

Genetic heterogeneity



Ingo Ruczinski

Assessing variants in the human genome

From the New Yorker



"O.K., let's slowly lower in the grant money." Todd Bearson Arlington, Mass.

NHLBI exome sequencing project



Program Officer: Deborah Applebaum-Bowden

Heart GO

PI: Stephen Rich University of Virginia

Lung GO

Pls: Michael Bamshad U. of Washington Kathleen Barnes, Johns Hopkins University

Women's Health Initiative Sequencing Program

PI: Rebecca Jackson Ohio State University



UNIVERSITY of WASHINGTON

Pls: Debbie Nickerson, Mark Rieder, Jay Shendure, & Phil Green



Pls: Stacey Gabriel & David Altshuler

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