Achieving Optimal Covariate Balance Under General Treatment Regimes

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Marc Ratkovic, Princeton University Optimal Covariate Balance

For many questions of interest in the social sciences, experiments are not possible \Rightarrow Possible bias in effect estimates

Regression adjustment or inverse weighting can be used to adjust for selection bias

 \Rightarrow Model dependence

Matching reduces bias and model dependence by identifying a set of untreated observations that are similar to the treated observations

Problems with Existing Matching Methods

Existing matching methods, such as propensity matching, Genetic matching, and Coarsened Exact Matching,

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- are sensitive to these choices
- have no formal statistical properties
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- can only handle a binary treatment can also accommodate continuous treatments

The Setup

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- Treatment: *T_i*
 - Binary treatment: $T_i \in \{0, 1\}$
 - Continuous treatment: $T_i \in (a, b)$
- Potential outcome: $Y_i(t)$
- Pre-treatment covariates: *X_i*
- IID observations $(Y_i(T_i), T_i, X_i)$ observed

Assumptions

- No interference among units
- Treatment occurs with uncertainty
- No omitted variables

Goal of Matching: Identify a subset of the data such that the covariates are balanced

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Common estimands identified on a balanced subset of the data:

- Average Treatment Effect: $E(Y_i(1) Y_i(0))$
- Average Treatment Effect on the Treated: $E(Y_i(1) - Y_i(0)|T_i = 1)$

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The logic of the proposed method proceeds in three steps:

- The optimality condition for an SVM sets an inner product between the treatment level and a covariate to zero
- Centering the treatment and covariate transforms this inner product to balance-in-mean or zero covariance.
- Balancing along a nonparametric basis extends the mean/covariance result to joint independence.

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Transform T_i from {0, 1} to {-1, 1}:

$$T_i^* = 2T_i - 1$$

Define the "hinge loss" $|z|_+ = \max(z, 0)$

Loss function:

$$\mathcal{L}(\boldsymbol{\beta}) = \sum_{i} |1 - T_{i}^{*} X_{i}^{*\top} \boldsymbol{\beta}|_{+} \quad \text{s.t. } X_{i}^{*^{\top}} \boldsymbol{\beta} \cdot \mathbf{1}(T_{i} = 1) < 1$$

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- $T_i^* = -1$; $X_i^{*\top}\beta = -0.5$: $|1 (-1) \cdot (-0.5)|_+ = |0.5|_+ = 0.5$ Hard to classify

The constraint keeps the loss for all treated observations as non-zero to identify the ATT.

Geometric Intuition of Proposed Method

Properly classified cases outside the margin are "easy-to-classify."

Cases in the margin, or improperly classified, have a treatment assignment estimated with some uncertainty.



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Law of Large Numbers gives

$$E(X_i|T_i=1) = E(X_i|T_i=0, i \in \mathcal{M})$$

The Binary Treatment SVM

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To achieve joint independence (**Proposition 1**):

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- Add a regularization term, to balance covariate imbalance and model complexity
- Observations in $\mathcal M$ are balanced

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The proof follows nearly exactly as the linear case, except in a high-dimensional space.

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Extension to a nonparametric function of X_i transforms uncorrelatedness to joint independence (**Proposition 2**).

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 - In a simple randomized experiment, the sample size is twice the expected misclassification loss
 - Number of balanced observations approaches twice the expected misclassification loss asymptotically
- Answers question of "how many matches"
- Tuning parameters selected through GACV criterion
- Identifies observations that appear to follow a simple randomization
 - Most useful when researcher does not know which variables to match finely, exactly, in mean, etc.

The 1975-1978 National Supported Work Study (Lalonde 1986)

- Treatment: job training, close management, peer support
- Recipients: welfare recipients, ex-addicts, young school dropouts, and ex-offenders
- *n*=445: 260 treated; 185 control
- PSID data used for matching, *n*=2490
- X: age, years of education, race, marriage status, high school degree, 1974 earnings, 1975 earnings, zero earnings in 1974, zero earnings in 1975

Analyses

Competitors

- Logistic propensity matching (Ho, et al. 2011)
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Datasets

- Experimental treated and untreated observations
- Experimental treated observations; observational untreated observations



Density of Treatment Effect Estimates Across Model Specifications, Using NSW Experimental Data



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Observational Results



Density of Treatment Effect Estimates Across Model Specifications, Untreated Observations Taken from Observational PSID Data

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Density of Treatment Effect Estimates Across Model Specifications, Untreated Observations Taken from Observational PSID Data

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Density of Treatment Effect Estimates Across Model Specifications, Untreated Observations Taken from Observational PSID Data The 1987 National Medical Expenditure Survey (Johnson, et al. 2003; Imai and van Dyk 2004)

- Treatment: *log(pack years*): packs a day times number of years smoking, logged
- Respondents: Representative sample of US population
- n = 9,708 smokers; to be balanced
- n = 9,804 non-smokers; reference group
- Outcome: Medical expenditure, dollars
- X: age at survey, age when started smoking, gender, race, education, marital status, census region, poverty status, seat-belt use

Assessing Balance



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Quantile Plot of Coefficient p-values from Regressing the Treatment On Pretreatment Covariates, Versus a Uniform Distribution



Estimated Effect



Medical Expenditures Relative to Non-Smokers

The proposed method adapts the SVM technology to the matching problem.

The method:

- is fully automated
- makes no functional form assumptions
- identifies the largest balanced subset
- can also accommodate continuous treatments